SUPPLEMENT TO "LEAVE-ONE-OUT SINGULAR SUBSPACE PERTURBATION ANALYSIS FOR SPECTRAL CLUSTERING"

BY Anderson Y. Zhang and Harrison H. Zhou

University of Pennsylvania and Yale University

APPENDIX A: PROOF OF THEOREM [2.3](#page-6-0)

The proof idea is similar to that of Theorem [2.2](#page-6-1) but with more involved calculation as *r* is not necessarily κ . Consider any $i \in [n]$. Define

$$
\tilde{\rho}_{-i}:=\frac{\hat{\lambda}_{-i,r}-\hat{\lambda}_{-i,r+1}}{\left\|\left(I-\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right)X_i\right\|}
$$

.

We need to verify $\tilde{\rho}_{-i} > 2$ first in order to apply Theorem [2.1.](#page-3-0) Recall the definition of P_{-i} in [\(36\)](#page--1-0) and E_{-i} in [\(38\)](#page--1-1). Let the SVD of P_{-i} be

$$
P_{-i} = \sum_{j=1}^{p \wedge (n-1)} \lambda_{-i,j} u_{-i,j} v_{-i,j}^T,
$$

where $\lambda_{-i,1} \geq \lambda_{-i,2} \geq \ldots \geq \lambda_{-i,p \wedge (n-1)}$. Denote $U_{-i,1}:r = (u_{-i,1}, u_{-i,2}, \ldots, u_{-i,r}) \in \mathbb{O}^{p \times r}$. Then by Weyl's inequality, we have

(48)
$$
|\hat{\lambda}_{-i,r} - \lambda_{-i,r}|, |\hat{\lambda}_{-i,r+1} - \lambda_{-i,r+1}| \leq ||E_{-i}|| \leq ||E||.
$$

Then the numerator

(49)
$$
\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1} \geq \lambda_{-i,r} - \lambda_{-i,r+1} - 2 ||E||.
$$

In the following, we are going to connect $\lambda_{-i,r} - \lambda_{-i,r+1}$ with $\lambda_r - \lambda_{r+1}$. To bridge the gap between $\lambda_{-i,r}$, $\lambda_{-i,r+1}$ and λ_r , λ_{r+1} , define

$$
\tilde{P}_{-i} := (\theta^*_{z^*_1}, \dots, \theta^*_{z^*_{i-1}}, U_{-i,1:r} U_{-i,1:r}^T \theta^*_{z^*_i}, \theta^*_{z^*_{i+1}}, \dots, \theta^*_{z^*_n}) \in \mathbb{R}^{p \times n}.
$$

Let $\tilde{\lambda}_{-i,1} \ge \tilde{\lambda}_{-i,2} \ge \ldots \ge \tilde{\lambda}_{-i,p}$ be its singular values. Note that $U_{-i,1:r} U_{-i,1:r}^T \tilde{P}_{-i}$ is the best rank-*r* approximation of \tilde{P}_{-i} . This is because for any rank-*r* projection matrix $M \in$ $\mathbb{R}^{p \times p}$ such that $M^2 = M$, we have

$$
\left\| \tilde{P}_{-i} - MM^T \tilde{P}_{-i} \right\|_{\mathrm{F}}^2 = \left\| (I - MM^T) P_{-i} \right\|_{\mathrm{F}}^2 + \left\| (I - MM^T) U_{-i, 1:r} U_{-i, 1:r}^T \theta_{z_i^*}^* \right\|_{\mathrm{F}}^2
$$

\n
$$
\geq \left\| (I - U_{-i, 1:r} U_{-i, 1:r}^T) P_{-i} \right\|_{\mathrm{F}}^2 + 0
$$

\n
$$
= \left\| \tilde{P}_{-i} - U_{-i, 1:r} U_{-i, 1:r}^T \tilde{P}_{-i} \right\|_{\mathrm{F}}^2,
$$

where we use the fact $U_{-i,1:r}U_{-i,1:r}^T P_{-i}$ is the best rank-*r* approximation of P_{-i} . Hence, span($U_{-i,1:r}$) is exactly the leading *r* left singular space of P_{-i} . It immediately implies:

• $\lambda_{-i,j} = \lambda_{-i,j}$ for any $j \geq r+1$, including

$$
\tilde{\lambda}_{-i,r+1} = \lambda_{-i,r+1}.
$$

• Since $U_{-i,1:r}U_{-i,1:r}^T \tilde{P}_{-i}$ and $U_{-i,1:r}U_{-i,1:r}^T P_{-i}$ only differ by one column where the latter one can be seen as the leave-one-out counterpart of the former one, using the same argument as in (37) , we have

(51)
$$
\lambda_{-i,r}^2 \ge \left(1 - \frac{k}{\beta n}\right) \tilde{\lambda}_{-i,r}^2.
$$

Then from [\(49\)](#page--1-3), we have

(52)
$$
\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1} \ge \sqrt{1 - \frac{k}{\beta n}} \tilde{\lambda}_{-i,r} - \tilde{\lambda}_{-i,r+1} - 2 \|E\|.
$$

For the difference between $\tilde{\lambda}_{i,r}$, $\tilde{\lambda}_{i,r+1}$ and λ_r , λ_{r+1} , we use the Weyl's inequality again:

$$
\max_{j \in [k]} \left| \tilde{\lambda}_{-i,j} - \lambda_j \right| \le \left\| P - \tilde{P}_{-i} \right\| = \left\| \theta_{z_i^*}^* - U_{-i,1:r} U_{-i,1:r}^T \theta_{z_i^*}^* \right\|.
$$

In the proof of Theorem [2.2,](#page-6-1) we show $u_{-i,j} \in \text{span}(\{\theta_a^*\}_{a \in [k]})$ for each $j \in [\kappa]$. Then

$$
\left\|\theta_{z_i^*}^* - U_{-i,1:r} U_{-i,1:r}^T \theta_{z_i^*}^*\right\| = \left\|\left(u_{-i,r+1}, \ldots, u_{-i,\kappa}\right) \left(u_{-i,r+1}, \ldots, u_{-i,\kappa}\right)^T \theta_{z_i^*}^*\right\|
$$

$$
= \sqrt{\sum_{a \in [\kappa]: a \ge r+1} \left(u_{-i,a}^T \theta_{z_i^*}^*\right)^2}.
$$

For any $a \in [\kappa]$ such $a \ge r + 1$, we have

$$
\left(u_{-i,a}^T \theta_{z_i^*}^*\right)^2 \le \frac{1}{\left|\left\{j \in [n] : z_j^* = z_i^*\right\}\right| - 1} \sum_{j \in [n]: j \ne i, z_j^* = z_i^*} \left(u_{-i,a}^T \theta_{z_j^*}^*\right)^2 \le \frac{1}{\frac{\beta n}{k} - 1} (u_{-i,a}^T P_{-i})^2
$$

$$
\le \frac{\lambda_{-i,a}^2}{\frac{\beta n}{k} - 1} \le \frac{\lambda_{-i,r+1}^2}{\frac{\beta n}{k} - 1}.
$$

Hence, we obtain $\|\theta_{z_i^*}^* - U_{-i,1:r} U_{-i,1:r}^T \theta_{z_i^*}^* \| \le \sqrt{\kappa} \lambda_{-i,a} / \sqrt{\beta n/k - 1}$ and consequently,

(53)
$$
\max_{j\in[k]} \left|\tilde{\lambda}_{-i,j} - \lambda_j\right| \leq \frac{\sqrt{\kappa}\lambda_{-i,r+1}}{\sqrt{\frac{\beta n}{k}-1}}.
$$

Then together with [\(50\)](#page--1-4), we have $|\lambda_{-i,r+1} - \lambda_{r+1}| \leq \sqrt{\kappa} \lambda_{-i,r+1}/\sqrt{\beta n/k - 1}$ and hence

(54)
$$
\lambda_{-i,r+1} \leq \frac{\lambda_{r+1}}{1 - \frac{\sqrt{\kappa}}{\sqrt{\frac{\beta_n}{\kappa} - 1}}}.
$$

Denote $d := \beta n / k$. With [\(52\)](#page--1-5), we have

$$
\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1} \ge \sqrt{\frac{d-1}{d}} \left(\lambda_r - \frac{\lambda_{-i,r+1}}{\sqrt{d-1}} \right) - \left(\lambda_{r+1} + \frac{\lambda_{-i,r+1}}{\sqrt{d-1}} \right) - 2 \Vert E \Vert
$$
\n
$$
\ge \sqrt{\frac{d-1}{d}} \lambda_r - \lambda_{r+1} \left(1 + \left(\frac{1}{\sqrt{d}} + \frac{1}{\sqrt{d-1}} \right) \frac{1}{1 - \frac{\sqrt{\kappa}}{\sqrt{d-1}}} \right) - 2 \Vert E \Vert
$$
\n
$$
\ge \sqrt{\frac{d-1}{d}} \left(\lambda_r - \lambda_{r+1} - \frac{4}{\sqrt{d}} \lambda_{r+1} \right) - 2 \Vert E \Vert
$$
\n(55)\n
$$
\ge \frac{3}{4} \left(\lambda_r - \lambda_{r+1} - \frac{4}{\sqrt{d}} \lambda_{r+1} \right) - 2 \Vert E \Vert,
$$

where in the last two inequalities we use the assumption that $d/k \ge 10$. As a consequence, we have

$$
\tilde{\rho}_{-i}\geq \frac{\hat{\lambda}_{-i,r}-\hat{\lambda}_{-i,r+1}}{\left\|\left(I-\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right)X_i\right\|}\geq \frac{\frac{3}{4}\left(\lambda_r-\lambda_{r+1}-\frac{4}{\sqrt{d}}\lambda_{r+1}\right)-2\left\|E\right\|}{\left\|\left(I-\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right)X_i\right\|}.
$$

Next, we are going to simplify the denominator of the above display. Using the orthogonality of the singular vectors, we have

$$
\begin{split}\n&\left\|\left(I-\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right)\theta_{z_i^*}^*\right\| \\
&\leq \left\|\left(I-\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right)\theta_{z_i^*}^*\right\| + \left\|\left(\hat{u}_{-i,r+1},\ldots,\hat{u}_{-i,\kappa}\right)\left(\hat{u}_{-i,r+1},\ldots,\hat{u}_{-i,\kappa}\right)^T\theta_{z_i^*}^*\right\| \\
&= \left\|\left(I-\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right)\theta_{z_i^*}^*\right\| + \sqrt{\sum_{j=r+1}^{\kappa} \left(\hat{u}_{-i,j}^T\theta_{z_i^*}^*\right)^2} \\
&\leq \frac{3\sqrt{\kappa}\|E\|}{\sqrt{\frac{\beta n}{k}-1}} + \sqrt{\sum_{j=r+1}^{\kappa} \left(\frac{\hat{\lambda}_{-i,j}}{\sqrt{\frac{\beta n}{k}-1}} + \frac{\|E\|}{\sqrt{\frac{\beta n}{k}-1}}\right)^2} \\
&\leq \frac{3\sqrt{\kappa}\|E\|}{\sqrt{\frac{\beta n}{k}-1}} + \sqrt{\kappa}\left(\frac{\hat{\lambda}_{-i,r+1}}{\sqrt{\frac{\beta n}{k}-1}} + \frac{\|E\|}{\sqrt{\frac{\beta n}{k}-1}}\right),\n\end{split}
$$

where the second to the inequality is due to (41) and (44) . By (54) and the Weyl's inequality, we have

$$
\hat{\lambda}_{-i,r+1} \leq \lambda_{-i,r+1} + ||E|| \leq \frac{1}{1 - \frac{\sqrt{\kappa}}{\sqrt{\frac{\beta n}{\kappa} - 1}}} \lambda_{r+1} + ||E||.
$$

Then, with the assumption $\beta n/k^2 \ge 10$, we have

$$
\left\| \left(I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) \theta_{z_i^*}^* \right\| \le \frac{3\sqrt{\kappa} \|E\|}{\sqrt{\frac{\beta n}{k} - 1}} + \sqrt{\kappa} \left(\frac{\lambda_{r+1}}{\sqrt{\frac{\beta n}{k} - 1} - \sqrt{\kappa}} + \frac{2\|E\|}{\sqrt{\frac{\beta n}{k} - 1}} \right)
$$

$$
\le \frac{\sqrt{k\kappa}}{\sqrt{\beta n}} (6\|E\| + 2\lambda_{r+1}).
$$

Hence,

$$
\left\| \left(I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) X_i \right\| \le \left\| \left(I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) \theta_{z_i^*}^* \right\| + \left\| \left(I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) \epsilon_i \right\|
$$

$$
\le \frac{\sqrt{k\kappa}}{\sqrt{\beta n}} (6 \left\| E \right\| + 2\lambda_{r+1}) + \left\| E \right\|.
$$

As a result,

$$
\tilde{\rho}_{-i}\geq\frac{\frac{3}{4}\left(\lambda_r-\lambda_{r+1}-\frac{4}{\sqrt{\beta n/k}}\lambda_{r+1}\right)-2\left\|E\right\|}{\frac{\sqrt{k\kappa}}{\sqrt{\beta n}}(6\left\|E\right\|+2\lambda_{r+1})+\left\|E\right\|}\geq\frac{\tilde{\rho}_0}{8}>2,
$$

under the assumption that $\beta n/(k^2) \ge 10$ and [\(11\)](#page-6-2).

The remaining part of the proof is to study $\{\hat{u}^T_{-i,a}X_i\}_{a\in[r]}$ and then apply Theorem [2.1.](#page-0-0) Following the exact argument as in the proof of Theorem $2.\overline{2}$, we have

$$
\sqrt{\sum_{a \in r} \left(\frac{\hat{u}^T_{-i,a}X_i}{\hat{\lambda}_{-i,a}}\right)^2} \leq \frac{\sqrt{r}}{\sqrt{\frac{\beta n}{k}-1}} + \frac{1}{\hat{\lambda}_{-i,r}} \frac{\|E\|\sqrt{r}}{\sqrt{\frac{\beta n}{k}-1}} + \frac{1}{\hat{\lambda}_{-i,r}}\left\|\hat{U}_{-i,1:r}\hat{U}^T_{-i,1:r}\epsilon_i\right\|.
$$

Under the assumption that $\beta n/(k^2) \ge 10$ and [\(11\)](#page-0-2), [\(55\)](#page-0-3) is lower bounded by $\lambda_r/2$. This also implies $\lambda_{-i,r} \geq \lambda_r/2$. Then a direct application of Theorem [2.1](#page-0-0) leads to

$$
\left\| \hat{U}_{1:r} \hat{U}_{1:r}^T - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right\|_{\rm F} \leq \frac{4\sqrt{2}}{\tilde{\rho}_{-i}} \left(\frac{\sqrt{r}}{\sqrt{\beta n/k - 1}} + \frac{1}{\hat{\lambda}_{-i,r}} \left(\frac{\sqrt{r} \left\| E \right\|}{\sqrt{\beta n/k - 1}} + \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\| \right) \right)
$$

$$
\leq \frac{128}{\tilde{\rho}_0} \left(\frac{\sqrt{kr}}{\sqrt{\beta n}} + \frac{\left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\|}{\lambda_r} \right).
$$

APPENDIX B: PROOFS OF RESULTS IN SECTION [3.4](#page-0-4)

Before presenting the proof of Lemma [3.3,](#page-0-5) we first show \hat{r} defined in [\(23\)](#page-0-6) always exists. In addition, since $\hat{r} \in [k]$ is a random variable, we are going to associate it with some deterministic set in [k]. Recall $\lambda_1 \geq \lambda_2 \geq \ldots \geq \lambda_{p \wedge n}$ are singular values of the signal matrix *P* and κ is the its rank. Let its SVD be $P = \sum_{i \in [p \wedge n]}^{\infty} \lambda_i u_i v_i^T$ with $\{u_j\}_{j \in [p \wedge n]} \in \mathbb{R}^p$ being its left singular vectors.

LEMMA B.1. *Under the same conditions as stated in Lemma [3.3,](#page-0-5) r*ˆ *always exists. Furthermore, we have* $\hat{r} \in \mathcal{R}$ *where*

(56)
$$
\mathcal{R} := \{a \in [k] : \lambda_a - \lambda_{a+1} \geq (\tilde{\rho} - 2) ||E|| \text{ and } \lambda_{a+1} \leq (k\tilde{\rho} + 1) ||E||\}.
$$

PROOF. The existence of \hat{r} can be proved by contradiction. If \hat{r} does not exist, it means that $\{a \in [k]: \lambda_a - \lambda_{a+1} \geq T\}$ is empty, which implies $\lambda_1 < \lambda_{k+1} + kT = \lambda_{k+1} + k\tilde{\rho}||E||$. By Weyl's inequality, we have $|\lambda_a - \lambda_a| \leq ||E||$ for all singular values of *X* and *P*. Then we have $\lambda_1 < (k\tilde{\rho} + 1) ||E||$. On the other hand, we have

$$
\lambda_1^2 = \max_{w \in \mathbb{R}^p : ||w|| = 1} ||w^T P||^2 \ge \max_{a, b \in [k]: a \ne b} \max_{w \in \mathbb{R}^p : ||w|| = 1} \frac{\beta n}{k} (||w^T \theta_a^*||^2 + ||w^T \theta_b^*||^2)
$$

$$
\ge \max_{a, b \in [k]: a \ne b} \max_{w \in \mathbb{R}^p : ||w|| = 1} \frac{\beta n}{2k} ||w^T \theta_a^* - w^T \theta_b^*||^2 = \frac{\beta n}{2k} \Delta^2,
$$

where the first inequality is due to the mixture model structure in *P* and the second inequality is due to $2(x_1 + x_2)^2 \ge (x_1 - x_2)^2$ for any two scalars x_1, x_2 . Then we have $\lambda_1 \geq \sqrt{\beta n/(2k)} \Delta = (\psi_0/\sqrt{2})k^{1.5} ||E||$ by [\(25\)](#page-0-7). Since $\tilde{\rho} < \tilde{\psi}_0/64$ is assumed, we have $(k\tilde{\rho}+1)\|E\|<(\tilde{\psi}_0/\sqrt{2})k^{1.5}$ $\|E\|$, which is a contradiction.

To prove the second statement, note that we have $\hat{\lambda}_{\hat{r}} - \hat{\lambda}_{\hat{r}+1} \ge \tilde{\rho} ||E||$ and $\hat{\lambda}_{\hat{r}+1} \le k\tilde{\rho} ||E||$. Since $|\hat{\lambda}_a - \lambda_a| \le ||E||$ for all singular values of *X* and *P*, we have $\lambda_{\hat{r}} - \lambda_{\hat{r}+1} \ge (\tilde{\rho} - 2) ||E||$
and $\lambda_{\hat{r}+1} \le (k\tilde{\rho}+1) ||E||$ Hence $\hat{r} \in \mathcal{R}$. and $\lambda_{\hat{r}+1} \leq (k\tilde{\rho}+1) ||E||$. Hence, $\hat{r} \in \mathcal{R}$.

PROOF OF LEMMA [3.3.](#page-0-5) From Lemma [B.1,](#page-0-8) we know \hat{r} exists and $\hat{r} \in \mathcal{R}$. Consider an arbitrary $r \in \mathcal{R}$ and define $\hat{U}_{1:r} := (\hat{u}_1, \dots, \hat{u}_r) \in \mathbb{R}^{p \times r}$. Perform *k*-means on the columns of $\hat{U}_{1:r} \hat{U}_{1:r}^T X$ and let the output be

$$
\left(\check{z}(r), \{\check{\theta}_j(r)\}_{j=1}^k\right) = \underset{z \in [k]^n, \{\theta_j\}_{j=1}^k \in \mathbb{R}^p}{\text{argmin}} \sum_{i \in [n]} \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T X - \theta_{z_i} \right\|^2.
$$

In the following, we are going to establish statistical properties for $\check{z}(r)$ and eventually obtain a desired upper bound for $\ell(\check{z}(r), z^*)$. Since performing *k*-means on the columns of $\hat{U}_{1:r}^T X$

is equivalent to *k*-means on the columns of $\hat{U}_{1:r}\hat{U}_{1:r}^T X$, and since $\hat{r} \in \mathcal{R}$, we have $\tilde{z} = \tilde{z}(\hat{r})$ and thus the desired upper bound also holds for $\ell(\tilde{z}, z^*)$.

In the rest of the proof we are going to analyze $\check{z}(r)$ for any $r \in \mathcal{R}$. For simplicity, we use the notation \tilde{z} , $\{\tilde{\theta}_j\}_{j\in[n]}$ instead of $\tilde{z}(r)$, $\{\tilde{\theta}_j(r)\}_{j\in[n]}$. The remaining proof can be decomposed into several parts.

(Preliminary Results for \check{z} , $\{\check{\theta}_j\}_{j \in [n]}$ *).* We are going to use Proposition [3.1](#page-0-9) to have some preliminary results. Define $U_{1:r} := (u_1, \ldots, u_r)$ and $U_{(r+1):k} := (u_{r+1}, \ldots, u_k)$. Instead of the decomposition (6) , we can write

$$
X_i = U_{1:r} U_{1:r}^T \theta_{z_i^*}^* + U_{(r+1):k} U_{(r+1):k}^T \theta_{z_i^*}^* + \epsilon_i = U_{1:r} U_{1:r}^T \theta_{z_i^*}^* + \check{\epsilon}_i,
$$

where $\check{\epsilon}_i := U_{(r+1):k} U_{(r+1):k}^T \theta_{z_i^*}^* + \epsilon_i$. In this way, we have a new mixture model with the centers being $\{U_{1:r}U_{1:r}^T \theta_a^*\}_{a \in [k]}$ and the additive noises being $\{\check{\epsilon}_i\}$. Define $\check{E} := (\check{\epsilon}_1, \dots, \check{\epsilon}_n)$. Then

$$
\|\check{E}\| \le \|E\| + \left\| \left(U_{(r+1):k} U_{(r+1):k}^T \theta_{z_1^*}^*, \dots, U_{(r+1):k} U_{(r+1):k}^T \theta_{z_n^*}^* \right) \right\|
$$

= \|E\| + \|U_{(r+1):k} U_{(r+1):k}^T P\| = \|E\| + \lambda_{r+1}
(57) $\le (k\tilde{\rho} + 2) \|E\|.$

The separation among the new centers is no longer Δ . Define

$$
\check{\Delta} := \min_{a,b \in [k]: a \neq b} \left\| U_{1:r} U_{1:r}^T \theta_a^* - U_{1:r} U_{1:r}^T \theta_b^* \right\|.
$$

For any $a, b \in [k]$, $U_{1:r}U_{1:r}^T \theta_a^* - U_{1:r}U_{1:r}^T \theta_b^* = (\theta_a^* - \theta_b^*) - U_{(r+1):k}U_{(r+1):k}^T \theta_a^* + U_{(r+1):k}U_{(r+1):k}^T \theta_b^*$. Also,

$$
\max_{a \in [k]} \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_a^* \right\| = \max_{a \in [k]} \sqrt{\frac{\sum_{i \in [n]: z_i^* = a} \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_a^* \right\|^2}{|\{i \in [n]: z_i^* = a\}|}} \le \frac{\left\| U_{(r+1):k} U_{(r+1):k}^T P \right\|_{\mathcal{F}}}{\sqrt{\beta n/k}}
$$
\n(58)\n
$$
\le \frac{2\sqrt{k}\lambda_{r+1}}{\sqrt{\beta n/k}} \le \frac{\sqrt{k}(k\tilde{\rho}+1) \left\| E \right\|}{\sqrt{\beta n/k}}.
$$

Hence, we have

$$
(59) \quad \check{\Delta} \ge \min_{a,b \in [k]: a \neq b} \|\theta_a^* - \theta_b^*\| - 2 \max_{a \in [k]} \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_a^* \right\| \ge \Delta - \frac{2\sqrt{k}(k\tilde{\rho}+1)\|E\|}{\sqrt{\beta n/k}}.
$$

Then from Proposition [3.1,](#page-0-9) as long as (which will be verified later)

(60)
$$
\check{\psi}_0 := \frac{\check{\Delta}}{\beta^{-0.5} k n^{-0.5} ||\check{E}||} \ge 16,
$$

we have

$$
\ell(\check{z}, z^*) = \frac{1}{n} |i \in [n] : \check{z}_i \neq \phi(z_i^*) \le \frac{C_0 k \left\| \check{E} \right\|^2}{n \check{\Delta}^2},
$$

and

$$
\max_{a \in [k]} \left\| \check{\theta}_{\phi(z)} - U_{1:r} U_{1:r}^T \theta_a^* \right\| \leq C_0 \beta^{-0.5} k n^{-0.5} \left\| \check{E} \right\|.
$$

where $C_0 = 128$.

(Entrywise Decomposition for \check{z} *).* Next, we are going to have an entrywise decomposition for $\mathbb{I}\{\hat{z}_i \neq \phi(z_i^*)\}$ that is analogous to that of Lemma [3.2.](#page-0-11) When [\(60\)](#page-0-12) is satisfied, from Lemma [3.1,](#page-0-13) we have

$$
\mathbb{I}\left\{\check{z}_i \neq \phi(z_i^*)\right\} \leq \mathbb{I}\left\{\left(1 - C_0\check{\psi}_0^{-1}\right)\check{\Delta} \leq 2\left\|\hat{U}_{1:r}\hat{U}_{1:r}^T\check{\epsilon}_i\right\|\right\}.
$$

By the definition of $\check{\epsilon}_i$ and [\(58\)](#page-0-14), we have

$$
\left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \check{\epsilon}_i \right\| \le \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| + \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T U_{(r+1):k} U_{(r+1):k}^T \theta_{z_i^*}^* \right\|
$$

$$
\le \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| + \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_{z_i^*}^* \right\|
$$

$$
\le \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| + \frac{\sqrt{k(k\tilde{\rho}+1) \|E\|}}{\sqrt{\beta n/k}}.
$$

Then, we have

$$
\mathbb{I}\left\{\check{z}_{i} \neq \phi(z_{i}^{*})\right\} \leq \mathbb{I}\left\{\left(1 - C_{0}\check{\psi}_{0}^{-1}\right)\check{\Delta} \leq 2\left(\left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}\epsilon_{i}\right\| + \frac{\sqrt{k}(k\tilde{\rho}+1)\left\|E\right\|}{\sqrt{\beta n/k}}\right)\right\}
$$
\n
$$
= \mathbb{I}\left\{\left(1 - C_{0}\check{\psi}_{0}^{-1} - \frac{2\sqrt{k}(k\tilde{\rho}+1)\left\|E\right\|}{\sqrt{\beta n/k}\check{\Delta}}\right)\check{\Delta} \leq 2\left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}\epsilon_{i}\right\|\right\}.
$$

From [\(56\)](#page-0-15), under the assumption that $\tilde{\rho} > 4$ and $\beta n / k^4 > 400$, we have $\tilde{\rho}_0$ defined as in [\(11\)](#page-0-2) to satisfy

$$
\tilde{\rho}_0 \ge \frac{(\tilde{\rho} - 1) \|E\|}{\max \left\{ \|E\| , \sqrt{\frac{k^2}{\beta n}} (k\tilde{\rho} + 1) \|E\| \right\}} \ge 2.
$$

Then Theorem [2.3](#page-0-16) can be applied, with which we have

$$
\left\|\hat{U}_{1:r}\hat{U}_{1:r}^T - \hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right\|_{\rm F} \le \frac{256\sqrt{rk}}{\sqrt{n\beta}} + \frac{256\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\epsilon_i\right\|}{\lambda_r}.
$$

Then following the proof of Lemma [3.2,](#page-0-11) we have

$$
\begin{split} & \mathbb{I}\left\{\check{z}_{i}\neq\phi(z_{i}^{*})\right\} \\ & \leq \mathbb{I}\left\{\left(1-C_{0}\check{\psi}_{0}^{-1}-\frac{2\sqrt{k}(k\tilde{\rho}+1)\left\|E\right\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|+\left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}-\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\right\|_{\mathrm{F}}\left\|E\right\|\right)\right\} \\ & \leq \mathbb{I}\left\{\left(1-C_{0}\check{\psi}_{0}^{-1}-\frac{2\sqrt{k}(k\tilde{\rho}+1)\left\|E\right\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(\frac{256\sqrt{rk}\left\|E\right\|}{\sqrt{n\beta}}+\left(1+\frac{256\left\|E\right\|}{\lambda_{r}}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right)\right\} \\ & \leq \mathbb{I}\left\{\left(1-C_{0}\check{\psi}_{0}^{-1}-\frac{2\sqrt{k}(k\tilde{\rho}+257)\left\|E\right\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(1+\frac{256\left\|E\right\|}{\lambda_{r}}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\| \right\} \\ & \leq \mathbb{I}\left\{\left(1-C_{0}\check{\psi}_{0}^{-1}-\frac{2\sqrt{k}(k\tilde{\rho}+257)\left\|E\right\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(1+\frac{256}{\tilde{\rho}-2}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\| \right\}, \end{split}
$$

where in the last inequality we use $\lambda_r \geq (\tilde{\rho} - 2) ||E|| > 0$ (as long as $\tilde{\rho} > 2$) from [\(56\)](#page-0-15).

The last step of the proof is to simplify the above display using Δ instead of Δ . Then, under the assumption that $\tilde{\rho} > 256$, we have $(1 + 256/(\tilde{\rho} - 2))^{-1} \le (1 - 512/\tilde{\rho})$. Recall the definition of ψ_0 in [\(25\)](#page-0-7). Under the assumption that $\tilde{\rho} \leq \psi_0/64$, we have

(61)
$$
\check{\Delta} \geq \Delta \left(1 - \frac{4\beta^{-0.5} k^2 n^{-0.5} \tilde{\rho} ||E||}{\Delta} \right) = \Delta \left(1 - \frac{4\tilde{\rho}}{\tilde{\psi}_0} \right) \geq \frac{\Delta}{2},
$$

according to (59) . Then together with (57) , we can verify (60) holds due to

$$
\check{\psi}_0 \ge \frac{\Delta/2}{\beta^{-0.5}kn^{-0.5}(k\tilde{\rho}+2)\|E\|} \ge \frac{\Delta}{4\beta^{-0.5}k^2n^{-0.5}\tilde{\rho}\|E\|} = \frac{\tilde{\psi}_0}{4\tilde{\rho}} \ge 16.
$$

Rearranging all the terms with the help of [\(61\)](#page-0-19), we can simplify $\mathbb{I}\{\check{z}_i \neq \phi(z_i^*)\}$ into

$$
\mathbb{I}\left\{\tilde{z}_{i} \neq \phi(z_{i}^{*})\right\}
$$
\n
$$
\leq \mathbb{I}\left\{\left(1 - 4C_{0}\tilde{\rho}\tilde{\psi}_{0} - \frac{4\beta^{-0.5}k^{2}n^{-0.5}\tilde{\rho}||E||}{\Delta/2}\right)\left(1 - \frac{256}{\tilde{\rho}}\right)\left(1 - \frac{4\tilde{\rho}}{\tilde{\psi}_{0}}\right)\Delta \leq 2\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\}
$$
\n
$$
\leq \mathbb{I}\left\{\left(1 - 5C_{0}\tilde{\rho}\tilde{\psi}_{0}^{-1} - 256\tilde{\rho}^{-1}\right)\Delta \leq 2\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\}.
$$

PROOF OF THEOREM [3.2.](#page-0-20) Recall the definition of $\mathcal F$ in [\(46\)](#page-0-21). Then if $\mathcal F$ holds, by appropriate choices of C_1 , C_2 , we can verify the assumptions needed in Lemma [3.3](#page-0-5) hold, which lead to

$$
\mathbb{I}\left\{\tilde{z}_i \neq \phi(z_i^*)\right\} \mathbb{I}\left\{\mathcal{F}\right\} \leq \mathbb{I}\left\{\left(1 - C''(\rho_2\psi_2^{-1} + \rho_2^{-1})\right)\Delta \leq 2\left\|\hat{U}_{-i,1:\hat{r}}\hat{U}_{-i,1:\hat{r}}^T\epsilon_i\right\|\right\} \mathbb{I}\left\{\mathcal{F}\right\},\
$$

for some constant $C'' > 0$. Though \hat{r} is random, the proof of Lemma [3.3](#page-0-5) shows that $\hat{r} \in \mathcal{R} \subset \mathcal{R}$ [*k*] where R is defined in [\(56\)](#page-0-15). Note that for any $r \in [k]$, we can follow the proof of Theorem [3.1](#page-0-22) to show

$$
\mathbb{E} \mathbb{I} \left\{ \left(1 - C''(\rho_2 \psi_2^{-1} + \rho_2^{-1}) \right) \Delta \le 2 \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\| \right\} \le \exp \left(- (1 - C'''(\rho_2 \psi_2^{-1} + \rho_2^{-1})) \frac{\Delta^2}{8\sigma^2} \right),
$$

for some constant $C''' > 0$. Hence, the same upper bound holds for $\mathbb{E}[\{(1 - C''(\rho_2 \psi_2^{-1} +$ φ_2^{-1})) $\Delta \leq 2\|\hat{U}_{-i,1:\hat{r}}\hat{U}_{-i,1:\hat{r}}^T\hat{\epsilon}_i\|\}.$ The rest of the proof follows that of Theorem [3.1](#page-0-22) and is omitted here.

APPENDIX C: PROOF OF THEOREM [3.3](#page-0-23)

Define $\mathcal{F} = \{ ||E|| \le \sqrt{2}(\sqrt{n} + \sqrt{p})\sigma \}$. Then by Lemma B.1 of [\[27\]](#page-0-24), we have $\mathbb{P}(\mathcal{F}) \ge$ $1 - e^{-0.08n}$. Then under the event *F*, the assumption [\(26\)](#page-0-21) implies [\(16\)](#page-0-25) holds, and hence [\(17\)](#page-0-2) and [\(18\)](#page-0-26) hold. For simplicity, and without loss of generality, we can let ϕ in [\(17\)](#page-0-2)-(18) to be the identity, and we get

$$
\ell(\hat{z}, z^*) = \frac{1}{n} |\{i \in [n] : \hat{z}_i \neq z_i^*\}| \leq \frac{C_0 k \left(1 + \sqrt{\frac{p}{n}}\right)^2 \sigma^2}{\Delta^2},
$$

and

$$
\max_{a \in [k]} \left\| \hat{\theta}_a - \theta_a^* \right\| \le C_0 \beta^{-0.5} k \left(1 + \sqrt{\frac{p}{n}} \right) \sigma,
$$

where $C_0 > 0$ is some constant.

Denote $\hat{P} = \hat{U}_{1:k}\hat{U}_{1:k}^T X$ and let $\hat{P}_{i,i}$ be its *i*th column so that $\hat{P}_{i,i} = \hat{U}_{1:k}\hat{U}_{1:k}^T X_i$. We define $r \in [k]$ as (with $\lambda_{k+1} := 0$)

(62)
$$
r = \max\left\{j \in [k]: \lambda_j - \lambda_{j+1} \geq \tau \sqrt{n+p\sigma}\right\},\,
$$

for a sequence $\tau \to \infty$ to be determined later. We note that if $\Delta/(k^{\frac{3}{2}} \tau \beta^{\frac{1}{2}} (1 + p/n)^{\frac{1}{2}} \sigma) \to$ ∞ , the set $\{j \in [k]: \lambda_j - \lambda_{j+1} \geq \tau \sqrt{n+p\sigma}\}$ is not empty. Otherwise, this would imply $\lambda_1 \leq$ $k\tau\sqrt{n+p\sigma}$ which would contradict with the fact $\lambda_1 \geq \sqrt{\beta n/k}\Delta/(2\sigma)$ (see Proposition A.1) of $[27]$). By the definition of *r* in (62) , we immediately have

(63)
$$
\lambda_r - \lambda_{r+1} \geq \tau \sqrt{n+p} \sigma,
$$

(64) and
$$
\lambda_{r+1} \leq k\tau \sqrt{n+p\sigma}
$$
.

We split $\hat{U}_{1:k}$ into $(\hat{U}_{1:r}, \hat{U}_{(r+1):k})$ where $\hat{U}_{1:r} := (\hat{u}_1, \dots, \hat{u}_r)$ and $\hat{U}_{(r+1):k} := (\hat{u}_{r+1}, \dots, \hat{u}_k)$. We decompose $\hat{P}_{\cdot,i} = \hat{P}_{\cdot,i}^{(1)} + \hat{P}_{\cdot,i}^{(2)}$, where $\hat{P}_{\cdot,i}^{(1)} := \hat{U}_{1:r} \hat{U}_{1:r}^T \hat{P}_{\cdot,i}$ and $\hat{P}_{\cdot,i}^{(2)} := \hat{U}_{(r+1):k} \hat{U}_{(r+1):k}^T \hat{P}_{\cdot,i}$. Similarly, for each $a \in [k]$, we decompose $\hat{\theta}_a = \hat{\theta}_a^{(1)} + \hat{\theta}_a^{(2)}$, where $\hat{\theta}_a^{(1)} := \hat{U}_{1:r} \hat{U}_{1:r}^T \hat{\theta}_a$ and $\hat{\theta}_a^{(2)} := \hat{U}_{(r+1):k} \hat{U}_{(r+1):k}^T \hat{\theta}_a$. Due to the orthogonality of $\{\hat{u}_l\}_{l \in [k]}$, we obtain that for any $i \in [n]$ and any $a \in [k]$ such that $a \neq z_i^*$,

$$
\begin{split} \mathbb{I}\left\{\hat{z}_i = a\right\} &\leq \mathbb{I}\left\{\left\|\hat{P}_{\cdot,i}^{(1)} + \hat{P}_{\cdot,i}^{(2)} - \hat{\theta}_a^{(1)} - \hat{\theta}_a^{(2)}\right\|^2 \leq \left\|\hat{P}_{\cdot,i}^{(1)} + \hat{P}_{\cdot,i}^{(2)} - \hat{\theta}_{z_i^*}^{(1)} - \hat{\theta}_{z_i^*}^{(2)}\right\|^2\right\} \\ & = \mathbb{I}\left\{2\left\langle\hat{P}_{\cdot,i}^{(1)} - \hat{\theta}_{z_i^*}^{(1)}, \hat{\theta}_{z_i^*}^{(1)} - \hat{\theta}_a^{(1)}\right\rangle + \left\|\hat{\theta}_{z_i^*}^{(1)} - \hat{\theta}_a^{(1)}\right\|^2 \leq 2\left\langle\hat{P}_{\cdot,i}^{(2)}, \hat{\theta}_a^{(2)} - \hat{\theta}_{z_i^*}^{(2)}\right\rangle - \left\|\hat{\theta}_a^{(2)}\right\|^2 + \left\|\hat{\theta}_{z_i^*}^{(2)}\right\|^2\right\} \end{split}
$$

We denote $\tau'' = o(1)$ to be another sequence which we will specify later. Then the above display can be decomposed and upper bounded by

$$
\begin{aligned} \mathbb{I}\left\{\hat{z}_i=a\right\}\leq&\mathbb{I}\left\{\left\|\hat{\theta}_{z_i^*}^{(1)}-\hat{\theta}_{a}^{(1)}\right\|-\frac{\tau''\Delta^2+\left\|\hat{\theta}_{z_i^*}^{(2)}\right\|^2}{\left\|\hat{\theta}_{z_i^*}^{(1)}-\hat{\theta}_{a}^{(1)}\right\|}\leq 2\left\|\hat{P}_{\cdot,i}^{(1)}-\hat{\theta}_{z_i^*}^{(1)}\right\| \right\}\\ &+\mathbb{I}\left\{\tau''\Delta^2\leq 2\left\langle\hat{P}_{\cdot,i}^{(2)},\hat{\theta}_{a}^{(2)}-\hat{\theta}_{z_i^*}^{(2)}\right\rangle\right\}=:A_{i,a}+B_{i,a}. \end{aligned}
$$

Then

$$
\mathbb{E}\ell(\hat{z},z^*) \leq \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}\left[\left\{\hat{z}_i = a\right\}\right]
$$
\n
$$
\leq \mathbb{P}\left(\mathcal{F}^{\complement}\right) + \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} A_{i,a} \mathbb{I}\left\{\mathcal{F}\right\} + \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} B_{i,a} \mathbb{I}\left\{\mathcal{F}\right\}.
$$

We are going to establish upper bounds first for $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} B_{i,a} \mathbb{I} \{ \mathcal{F} \}$ and then for $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} A_{i,a} \mathbb{I} \{ \mathcal{F} \}.$

(Analysis on $n^{-1} \sum_{i \in [n]} \sum_{a \neq z_i^*} \mathbb{E} B_{i,a} \mathbb{I} \{\mathcal{F}\}\)$. For $\sum_{i \in [n]} \sum_{a \neq z_i^*} \mathbb{E} B_{i,a} \mathbb{I} \{\mathcal{F}\}\)$, we can di-

rectly use upper bounds established in Section 4.4.3 of $[27]$ ¹. It proves that for any $i \in [n]$,

$$
\sum_{a\in[k]:a\neq z_i^*}B_{i,a}\mathbb{I}\left\{\mathcal{F}\cap\mathcal{T}\right\}\leq 2\exp\left(-\frac{1}{2}\left(c_4\frac{\tau''\Delta}{k^{\frac{7}{2}}\tau^2\beta^{-\frac{1}{2}}(1+\frac{p}{n})\sigma}\sqrt{\frac{n-k}{3n}}\right)^2\frac{\Delta^2}{\sigma^2}\right),
$$

where $c_4 > 0$ is some constant, and $\mathcal T$ is some high-probability event in the sense that

$$
\mathbb{P}(\mathcal{T}) \ge 1 - nk \exp\left(-\frac{(n-k)}{9}\right).
$$

Hence,

$$
\frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} B_{i,a} \mathbb{I} \left\{ \mathcal{F} \right\} \leq \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} B_{i,a} \mathbb{I} \left\{ \mathcal{F} \cap \mathcal{T} \right\} + \mathbb{P} \left(\mathcal{T}^{\complement} \right)
$$
\n
$$
\leq 2 \exp \left(-\frac{1}{2} \left(c_4 \frac{\tau'' \Delta}{k^{\frac{7}{2}} \tau^2 \beta^{-\frac{1}{2}} (1 + \frac{p}{n}) \sigma} \sqrt{\frac{n-k}{3n}} \right)^2 \frac{\Delta^2}{\sigma^2} \right) + nk \exp \left(-\frac{(n-k)}{9} \right).
$$

(Analysis on $n^{-1} \sum_{i \in [n]} \sum_{a \neq z_i^*} \mathbb{E} A_{i,a} \mathbb{I} \{ \mathcal{F} \}$ *).* We first follow some algebra as in Section 4.4.2 of [\[27\]](#page-0-24) to simplify $A_{i,a} \mathbb{I} \{ \mathcal{F} \}$. For any $i \in [n]$ and $a \neq z_i^*$, it proves

(66)
$$
A_{i,a}\mathbb{I}\left\{\mathcal{F}\right\} \leq \mathbb{I}\left\{\left(1-c_1\tau''-\frac{c_1k^2\tau\beta^{-\frac{1}{2}}\sqrt{1+\frac{p}{n}}\sigma}{\Delta}\right)\Delta\leq 2\left\|\hat{P}_{\cdot,i}^{(1)}-\hat{\theta}_{z_i^*}^{(1)}\right\|\right\}\mathbb{I}\left\{\mathcal{F}\right\},\right\}
$$

for some constant $c_1 > 0$. Still working on the event \mathcal{F} , it also proves

(67)
$$
\left\| \hat{P}_{\cdot,i}^{(1)} - \hat{\theta}_{z_i^*}^{(1)} \right\| \leq \left\| \hat{P}_{\cdot,i}^{(1)} - \hat{U}_{1:r} \hat{U}_{1:r}^T \theta_{z_i^*}^* \right\| + 8\sqrt{2} \sqrt{\beta^{-1} k^2 \left(1 + \frac{p}{n} \right)} \sigma.
$$

Our following analysis on $A_{i,a}$ ^{$[\mathcal{F}]$} is different from the rest proof in Section 4.4.2 of [\[27\]](#page-0-24). Note that $\hat{P}_{i,i}^{(1)} - \hat{U}_{1:r} \hat{U}_{1:r}^T \theta_{z_i^*}^* = \hat{U}_{1:r} \hat{U}_{1:r}^T X_i - \hat{U}_{1:r} \hat{U}_{1:r}^T \theta_{z_i^*}^* = \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i$. Then [\(66\)](#page-0-29) and (67) give

$$
(68)\quad A_{i,a}\mathbb{I}\left\{\mathcal{F}\right\} \leq \mathbb{I}\left\{\left(1 - c_2\tau'' - \frac{c_2k^2\tau\beta^{-\frac{1}{2}}\left(1 + \sqrt{\frac{p}{n}}\right)\sigma}{\Delta}\right)\Delta \leq 2\left\|\hat{U}_{1:r}\hat{U}_{1:r}^T\epsilon_i\right\|\right\}\mathbb{I}\left\{\mathcal{F}\right\},\
$$

where we use $\tau \to \infty$ and the fact that $1 + \sqrt{p/n}$, $\sqrt{1 + p/n}$ are of the same order.

Recall the definition of X_{-i} in [\(8\)](#page-0-31) and $\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T$ is the leave-one-out counterpart of $\hat{U}_{1:r}\hat{U}_{1:r}^T$. For [\(68\)](#page-0-32), we can decompose $\|\hat{U}_{1:r}\hat{U}_{1:r}^T\epsilon_i\|$ into

$$
\left\|\hat{U}_{1:r}\hat{U}_{1:r}^T\epsilon_i\right\| \le \left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\epsilon_i\right\| + \left\|\hat{U}_{1:r}\hat{U}_{1:r}^T - \hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right\|_{\rm F}\left\|\epsilon_i\right\|.
$$

To upper bound $\|\hat{U}_{1:r}\hat{U}_{1:r}^T - \hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\|_F$, we are going to use Theorem [2.3.](#page-0-16) Since [\(63\)](#page-0-33)-[\(64\)](#page-0-34) hold, under the assumption $\beta n/k^4 \ge 100$, we have

$$
\frac{\lambda_r - \lambda_{r+1}}{\max\left\{\|E\|,\sqrt{\frac{k^2}{n\beta}}\lambda_{r+1}\right\}} \ge \frac{\tau}{2}.
$$

¹The model in [\[27\]](#page-0-24) assumes $\{\epsilon_j\} \stackrel{iid}{\sim} \mathcal{N}(0, I)$ while in this paper we assume $\{\epsilon_j\} \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2 I)$. To directly use results from [\[27\]](#page-0-24), we can re-scale our data to have $X'_j = X_j/\sigma$ for all $j \in [n]$. Then $\{X'_j\}$ has $\mathcal{N}(0, I)$ noise and the separation between their centers becomes Δ/σ . Then all the results from [\[27\]](#page-0-24) can be used here with Δ replaced by Δ/σ .

Applying Theorem [2.3,](#page-0-16) we have

$$
\left\|\hat{U}_{1:r}\hat{U}_{1:r}^T - \hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\right\|_{\rm F} \le \frac{256\sqrt{rk}}{\sqrt{n\beta}} + \frac{256\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\epsilon_i\right\|}{\lambda_r}.
$$

Hence,

$$
\left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| \le \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\| + \left(\frac{256\sqrt{rk}}{\sqrt{n\beta}} + \frac{256 \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\|}{\lambda_r} \right) \left\| E \right\|
$$

\n
$$
= \frac{256k \left\| E \right\|}{\sqrt{n\beta}} + \left(1 + \frac{256 \left\| E \right\|}{\lambda_r} \right) \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\|
$$

\n
$$
\le \frac{256\sqrt{2}k(\sqrt{n} + \sqrt{p})\sigma}{\sqrt{n\beta}} + \left(1 + \frac{256\sqrt{2}(\sqrt{n} + \sqrt{p})\sigma}{\tau \sqrt{n + p}\sigma} \right) \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\|
$$

\n
$$
\le 512k\beta^{-0.5} \left(1 + \sqrt{\frac{p}{n}} \right) \sigma + \left(1 + 512\tau^{-1} \right) \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\|,
$$

where in the second to the last inequality, we use [\(63\)](#page-0-33) for λ_r and the event *F* for *F*. Then [\(68\)](#page-0-32) leads to

$$
A_{i,a}\mathbb{I}\left\{\mathcal{F}\right\} \leq \mathbb{I}\left\{\left(1 - c_3\tau'' - \frac{c_3k^2\tau\beta^{-\frac{1}{2}}\left(1 + \sqrt{\frac{p}{n}}\right)\sigma}{\Delta}\right)\Delta \leq 2\left(1 + 512\tau^{-1}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\epsilon_i\right\|\right\}\mathbb{I}\left\{\mathcal{F}\right\}
$$

$$
\leq \mathbb{I}\left\{\left(1 - c_4\left(\frac{k^2\tau\beta^{-\frac{1}{2}}\left(1 + \sqrt{\frac{p}{n}}\right)\sigma}{\Delta} + \tau^{-1}\right)\right)\Delta \leq 2\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\epsilon_i\right\|\right\},\
$$

where $c_3, c_4 > 0$ are some constants. As long as $1 - c_4(k^2 \tau \beta^{-0.5}(1 + \sqrt{p/n})\sigma/\Delta + \tau^{-1}) >$ 1/2, we can use Lemma [E.2](#page-0-35) to calculate the tail probability of $\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\epsilon_i\|$. Following the proof of Theorem [3.1,](#page-0-22) we have

$$
\mathbb{E} A_{i,a} \mathbb{I} \left\{ \mathcal{F} \right\} \le \exp \left(- \left(1 - c_5 \left(\frac{k^2 \tau \beta^{-\frac{1}{2}} \left(1 + \sqrt{\frac{p}{n}} \right) \sigma}{\Delta} + \tau^{-1} \right) \right) \frac{\Delta^2}{8\sigma^2} \right),
$$

for some constant $c_5 > 0$. Then we have,

$$
n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} A_{i,a} \mathbb{I} \{\mathcal{F}\} \leq k \exp \left(-\left(1 - c_5 \left(\frac{k^2 \tau \beta^{-\frac{1}{2}} \left(1 + \sqrt{\frac{p}{n}}\right) \sigma}{\Delta} + \tau^{-1}\right)\right) \frac{\Delta^2}{8\sigma^2}\right).
$$

(*Obtaining the Final Result.*) From [\(65\)](#page-0-36) and the above upper bounds on $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} B_{i,a} \mathbb{I} \{ \mathcal{F} \}$ and $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E} A_{i,a} \mathbb{I} \{\mathcal{F}\}\)$, we have

$$
\mathbb{E}\ell(\hat{z},z^*) \le e^{-0.08n} + 2\exp\left(-\frac{1}{2}\left(c_4 \frac{\tau''\Delta}{k^{\frac{7}{2}}\tau^2\beta^{-\frac{1}{2}}(1+\frac{p}{n})\sigma}\sqrt{\frac{n-k}{3n}}\right)^2\frac{\Delta^2}{\sigma^2}\right) + nk\exp\left(-\frac{(n-k)}{9}\right) + k\exp\left(-\left(1 - c_5\left(\frac{k^2\tau\beta^{-\frac{1}{2}}\left(1+\sqrt{\frac{p}{n}}\right)\sigma}{\Delta} + \tau^{-1}\right)\right)\frac{\Delta^2}{8\sigma^2}\right).
$$

Since we assume $\beta n/k^4 \ge 100$, we have $(n - k)/n > 0.99$. Hence, under the assumption that $\Delta/(k^{3.5}\beta^{-0.5}(1+\frac{p}{n})\sigma) \to \infty$, we can take τ, τ'' to be

$$
\tau = \tau''^{-1} := \left(\frac{\Delta}{k^{3.5}\beta^{-0.5}\left(1 + \frac{p}{n}\right)\sigma}\right)^{0.25}
$$

such that $\tau \to \infty$ and $\tau'' = o(1)$. Then for some constant $c_6 > 0$, we have

$$
\mathbb{E}\ell(\hat{z},z^*) \le e^{-0.08n} + 2\exp\left(-\frac{c_4^2}{12}\left(\frac{\Delta}{k^{3.5}\beta^{-0.5}\left(1+\frac{p}{n}\right)\sigma}\right)^{0.5}\frac{\Delta^2}{\sigma^2}\right) + nke^{-0.1n}
$$

$$
+ k\exp\left(-\left(1 - 2c_5\left(\frac{\Delta}{k^{3.5}\beta^{-0.5}\left(1+\frac{p}{n}\right)\sigma}\right)^{-0.25}\right)\frac{\Delta^2}{8\sigma^2}\right)
$$

$$
\le \exp\left(-\left(1 - c_6\left(\frac{\Delta}{k^{3.5}\beta^{-0.5}\left(1+\frac{p}{n}\right)\sigma}\right)^{-0.25}\right)\frac{\Delta^2}{8\sigma^2}\right) + 2e^{-0.08n}.
$$

APPENDIX D: PROOFS OF RESULTS IN SECTION [3.6](#page-0-37)

D.1. Proof of Theorem [3.4.](#page-0-9) The proof of Theorem [3.4](#page-0-9) relies on the following entrywise decomposition that is analogous to Lemma [3.2](#page-0-11) but in an opposite direction. Note the the singular vectors \hat{u}_1 , and $\{\hat{u}_1, \hat{u}_2\}_{i \in [n]}$ are all identifiable up to sign. Without loss of generality, we assume $\langle \hat{u}_1, u_1 \rangle \ge 0$ and $\langle \hat{u}_1, -i, u_1 \rangle \ge 0$ for all $i \in [n]$.

LEMMA D.1. *Consider the model [\(28\)](#page-0-38). Let* $\phi \in \Phi$ *be the permutation such that* $\ell(\check{z},z^*) =$ $\frac{1}{n}$ $|\{i \in [n]: \check{z}_i \neq \phi(z_i^*)\}|$. Then there exists a constants $C, C_1 > 0$ such that if

(69)
$$
\frac{\Delta}{\beta^{-0.5}n^{-0.5}\|E\|} \geq C,
$$

then for any $i \in [n]$ *,*

$$
(70)\quad \mathbb{I}\left\{\check{z}_i \neq \phi(z_i^*)\right\} \ge \mathbb{I}\left\{\left(1 + \frac{C_1\beta^{-0.5}n^{-0.5}\|E\|}{\Delta}\right)\Delta \le -2(\hat{u}_{1, -i}^T\epsilon_i)sign(u_1^T\theta_{\phi(z_i^*)})\right\}.
$$

PROOF. The proof mainly follows the proofs of Lemma [3.1](#page-0-13) and Lemma [3.2](#page-0-11) with some modifications such as adding a negative term instead of a positive term in order to obtain a lower bound.

We first write \check{z} equivalently as

$$
\left(\check{z}, \{\check{\theta}_j\}_{j=1}^2\right) = \underset{z \in [2]^n, \{\theta_j\}_{j=1}^2 \in \mathbb{R}^p}{\text{argmin}} \sum_{i \in [n]} \left\|\hat{u}_1\hat{u}_1^T X_i - \theta_{z_i}\right\|^2,
$$

where $\hat{\theta}_a = \hat{u}_1 \check{c}_a$ for each $a \in [2]$. Note that $k = 2$. From Proposition [3.1,](#page-0-9) we have

$$
\frac{1}{n} |\{i \in [n] : \check{z}_i \neq \phi(z_i^*)\}| \le \frac{C_0 k ||E||^2}{n \Delta^2},
$$

and

(71)
$$
\max_{a \in [2]} \left\| \check{\theta}_{\phi(a)} - \theta_a^* \right\| \leq C_0 \beta^{-0.5} k n^{-0.5} \left\| E \right\|,
$$

for some permutation $\phi : [2] \rightarrow [2]$ and some constant $C_0 > 0$. Without loss of generality, assume $\phi =$ Id.

Recall that $\theta_1^* = -\theta_2^* = \delta 1_p$, $u_1 = 1/\sqrt{p}1_p$, $\lambda_1 = \delta \sqrt{np} = \frac{\Delta \sqrt{n}}{2}$, and $|u_1^T(\theta_{z_i^*}^* (-\theta_{z_i^*}^*)$) $| = 2\delta\sqrt{p} = \Delta$. By Davis-Kahan Theorem, we have

$$
\min_{s \in \pm 1} \|\hat{u}_1 - su_1\| \le \frac{\|E\|}{\lambda_1} = \frac{2\|E\|}{\sqrt{n}\Delta} \le 1/16,
$$

where the last inequality is due to the assumption [\(16\)](#page-0-25). Since we assume $\langle \hat{u}_1, u_1 \rangle \ge 0$, we have $\|\hat{u}_1 - s u_1\| = \min_{s \in +1} \|\hat{u}_1 - s u_1\|.$

Consider any $i \in [n]$ and any $a \in [2]$ such that $a \neq z_i^*$. Note that for any scalars x, y, w , if $|x - y| \le |x - w|$, we have equivalently sign $(w - y)(y + w)/2 \ge \text{sign}(w - y)x$. Since $(y + w)/2 = (y - w)/2 + w$, a sufficient condition is $|w - y|/2 + |w| \leq (-\text{sign}(w - y))x$. Hence, we have

$$
\begin{split}\n\mathbb{I}\left\{\left\|\hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{a}\right\| \leq\left\|\hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{z_{i}^{*}}\right\|\right\} \\
&= \mathbb{I}\left\{\left|\hat{u}_{1}^{T}X_{i}-\hat{u}_{1}^{T}\check{\theta}_{a}\right| \leq\left|\hat{u}_{1}^{T}X_{i}-\hat{u}_{1}^{T}\check{\theta}_{z_{i}^{*}}\right|\right\} \\
&= \mathbb{I}\left\{\left|\hat{u}_{1}^{T}\epsilon_{i}-\hat{u}_{1}^{T}\left(\check{\theta}_{a}-\theta_{z_{i}^{*}}^{*}\right)\right| \leq\left|\hat{u}_{1}^{T}\epsilon_{i}-\hat{u}_{1}^{T}\left(\check{\theta}_{z_{i}^{*}}-\theta_{z_{i}^{*}}^{*}\right)\right|\right\} \\
&\geq \mathbb{I}\left\{\frac{1}{2}\left|\hat{u}_{1}^{T}(\check{\theta}_{z_{i}^{*}}-\check{\theta}_{a})\right|+\left|\hat{u}_{1}^{T}\left(\check{\theta}_{z_{i}^{*}}-\theta_{z_{i}^{*}}^{*}\right)\right| \leq-(\hat{u}_{1}^{T}\epsilon_{i})\text{sign}(\hat{u}_{1}^{T}(\check{\theta}_{z_{i}^{*}}-\check{\theta}_{a}))\right\} \\
&\geq \mathbb{I}\left\{\left\|\check{\theta}_{z_{i}^{*}}-\check{\theta}_{a}\right\|+2\left\|\check{\theta}_{z_{i}^{*}}-\theta_{z_{i}^{*}}^{*}\right\| \leq-2(\hat{u}_{1}^{T}\epsilon_{i})\text{sign}(\hat{u}_{1}^{T}(\check{\theta}_{z_{i}^{*}}-\check{\theta}_{a}))\right\}.\n\end{split}
$$

We are going to show $sign(\hat{u}_1^T(\check{\theta}_{z_i^*} - \check{\theta}_a)) = sign(u_1^T(\theta_{z_i^*}^* - \theta_a^*))$. By [\(71\)](#page-0-39), we have

$$
\left\langle \check{\theta}_{z_i^*} - \check{\theta}_a, \theta_{z_i^*}^* - \theta_a^* \right\rangle = \left\| \theta_{z_i^*}^* - \theta_a^* \right\|^2 + \left\langle \check{\theta}_{z_i^*} - \theta_{z_i^*}^*, \theta_{z_i^*}^* - \theta_a^* \right\rangle + \left\langle \check{\theta}_a - \theta_a^*, \theta_{z_i^*}^* - \theta_a^* \right\rangle
$$

\n
$$
\geq \Delta^2 \left(1 - \frac{2C_0 k \beta^{-0.5} n^{-0.5} ||E||}{\Delta} \right)
$$

\n
$$
> 0,
$$

where the last inequality holds as long as $\Delta > 2C_0\beta^{-0.5}kn^{-0.5}$ ||E||. Due to the fact $\theta_{z_i^*}^* - \theta_a^* \in \text{span}(u_1)$, $\check{\theta}_{z_i^*} - \check{\theta}_a^* \in \text{span}(\hat{u}_1)$, and $\langle \hat{u}_1, u_1 \rangle \geq 0$, if $u_1, \theta_{z_i^*}^* - \theta_a^*$ are in the same direction, then $\hat{u}_1, \check{\theta}_{z_i^*} - \check{\theta}_a^*$ must also be in the same direction, and vice versa. Hence, we have $sign(\hat{u}_1^T(\check{\theta}_{z_i^*} - \check{\theta}_a)) = sign(u_1^T(\theta_{z_i^*}^* - \theta_a^*))$. Thus,

$$
\mathbb{I}\left\{\left\|\hat{u}_1\hat{u}_1^T X_i - \check{\theta}_a\right\| \leq \left\|\hat{u}_1\hat{u}_1^T X_i - \check{\theta}_{z_i^*}\right\|\right\} \n\geq \mathbb{I}\left\{\left\|\check{\theta}_{z_i^*} - \check{\theta}_a\right\| + 2\left\|\check{\theta}_{z_i^*} - \theta_{z_i^*}^*\right\| \leq -2(\hat{u}_1^T \epsilon_i) \text{sign}(u_1^T (\theta_{z_i^*}^* - \theta_a^*))\right\}.
$$

Following the same analysis as in the proof of Lemma [3.1,](#page-0-13) we can get the following result that is analogous to (45) :

$$
\mathbb{I}\left\{\left\|\hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{a}\right\| \leq\left\|\hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{z_{i}^{*}}\right\|\right\} \n\geq \mathbb{I}\left\{\left(1+\frac{4C_{0}\beta^{-0.5}kn^{-0.5}\left\|E\right\|}{\Delta}\right)\Delta \leq -2(\hat{u}_{1}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*}))\right\}.
$$

Next, we are going to decompose $\hat{u}_1^T \epsilon_i$ following the proof of Lemma [3.2.](#page-0-11) Denote $\hat{u}_{1,-i}$ be the leave-one-out counterpart of \hat{u}_1 , i.e., $\hat{u}_{1,-i}$ is the leading left singular vector of X_{-i} . Since we assume $\langle \hat{u}_{1,-i}, u_1 \rangle \ge 0$, we have $\|\hat{u}_{1,-i} - u_1\| \le 2 \|E\| / (\sqrt{n-1}\Delta)$. As a result, we have $\|\hat{u}_{1,-i} - \hat{u}_1\| \leq 4 \|\hat{E}\| / (\sqrt{n-1}\Delta)$ which leads to

(72)
$$
\langle \hat{u}_{1,-i}, \hat{u}_1 \rangle \ge 1 - 4 ||E|| / (\sqrt{n-1}\Delta) > 0.
$$

We have the following decomposition:

$$
\begin{split}\n& (\hat{u}_{1}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\
&= \langle \hat{u}_{1}, \hat{u}_{1}\hat{u}_{1}^{T}\epsilon_{i} \rangle \text{ sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\
&= \langle \hat{u}_{1}, (\hat{u}_{1,-i}\hat{u}_{1,-i}^{T})\epsilon_{i} \rangle \text{ sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \langle \hat{u}_{1}, (\hat{u}_{1}\hat{u}_{1}^{T}-\hat{u}_{1,-i}\hat{u}_{1,-i}^{T})\epsilon_{i} \rangle \text{ sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\
&= \langle \hat{u}_{1}, \hat{u}_{1,-i} \rangle (\hat{u}_{1,-i}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \langle \hat{u}_{1}, (\hat{u}_{1}\hat{u}_{1}^{T}-\hat{u}_{1,-i}\hat{u}_{1,-i}^{T})\epsilon_{i} \rangle \text{ sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\
&\leq \langle \hat{u}_{1}, \hat{u}_{1,-i} \rangle (\hat{u}_{1,-i}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + ||\hat{u}_{1}\hat{u}_{1}^{T}-\hat{u}_{1,-i}\hat{u}_{1,-i}^{T}|| \,||\epsilon_{i}|| \,.\n\end{split}
$$

Note that $\lambda_1 / ||E|| = \Delta \sqrt{n}/(2 ||E||)$ is greater than 16 under the assumption [\(69\)](#page-0-41) holds for a large constant *C*. From Theorem [2.2](#page-0-1) we have \mathbf{r} \sim

$$
\left\|\hat{u}_1\hat{u}_1^T - \hat{u}_{1,-i}\hat{u}_{1,-i}^T\right\| \le \frac{128}{\lambda_1/\left\|E\right\|} \left(\frac{k}{\sqrt{\beta n}} + \frac{\left\|\hat{u}_{1,-i}\hat{u}_{1,-i}^T\epsilon_i\right\|}{\lambda_1}\right).
$$

Then,

$$
\begin{split} &(\hat{u}_{1}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\ &\leq \langle \hat{u}_{1},\hat{u}_{1,-i}\rangle\,(\hat{u}_{1,-i}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \left(\frac{128k}{\sqrt{n\beta(\lambda_{1}/\left\Vert E\right\Vert)}}+\frac{128\left\Vert \hat{u}_{1,-i}\hat{u}_{1,-i}^{T}\epsilon_{i}\right\Vert }{\lambda_{1}^{2}/\left\Vert E\right\Vert }\right)\,\Vert E\Vert \\ & = \langle \hat{u}_{1},\hat{u}_{1,-i}\rangle\,(\hat{u}_{1,-i}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \frac{256n^{-0.5}k\beta^{-0.5}\left\Vert E\right\Vert^{2}}{\Delta}+\frac{512\left\vert \hat{u}_{1,-i}^{T}\epsilon_{i}\right\vert n^{-1}\left\Vert E\right\Vert^{2}}{\Delta^{2}}. \\ & \text{So far we have obtained} \\ & \mathbb{I}\left\lbrace \left\Vert \hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{a}\right\Vert \leq \left\Vert \hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{z_{i}^{*}}\right\Vert \right\rbrace \\ & \geq \mathbb{I}\left\lbrace \left(1+\frac{4C_{0}\beta^{-0.5}kn^{-0.5}\left\Vert E\right\Vert }{\Delta}\right)\Delta \leq -2\left\langle \hat{u}_{1},\hat{u}_{1,-i}\right\rangle\left\langle \hat{u}_{1,-i}^{T}\epsilon_{i}\right\rangle \text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \right. \\ & \left. -\frac{256n^{-0.5}k\beta^{-0.5}\left\Vert E\right\Vert^{2}}{\Delta}-\frac{512\left\vert \hat{u}_{1,-i}^{T}\epsilon_{i}\right\vert n^{-1}\left\Vert E\right\Vert^{2}}{\Delta^{2}}\right\rbrace \right\rbrace \\ & = \mathbb{I}\left\lbrace \left(1+\frac{4C_{0}\beta^{-0.5}kn^{-0.5}\left\Vert E
$$

From (72) we have

$$
\langle \hat{u}_{1,-i}, \hat{u}_1 \rangle - \frac{512n^{-1} ||E||^2}{\Delta^2} \ge 1 - 4 \frac{||E|| (n-1)^{-0.5}}{\Delta} - \frac{512n^{-1} ||E||^2}{\Delta^2}
$$

$$
\ge 1 - \frac{16n^{-0.5} ||E||}{\Delta} \ge \frac{1}{2},
$$

assuming $\frac{\Delta}{n^{-0.5}||E||} \ge 64$. For any $x, y, z, w \in \mathbb{R}$ such that $x \ge 0, 1 \ge z \ge 0$, and $z|y| > w \ge 0$, we have $\lim_{x \to a} \{x \leq z \leq y - w\} \geq \lim_{x \to a} \{x \leq (z - w/|y|) y\}$. We then have,

$$
\mathbb{I}\left\{\left\|\hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{a}\right\| \leq\left\|\hat{u}_{1}\hat{u}_{1}^{T}X_{i}-\check{\theta}_{z_{i}^{*}}\right\|\right\}
$$
\n
$$
\geq \mathbb{I}\left(\left(1+\frac{4C_{0}\beta^{-0.5}kn^{-0.5}\left\|E\right\|}{\Delta}+\frac{256n^{-0.5}k\beta^{-0.5}\left\|E\right\|^{2}}{\Delta^{2}}\right)\Delta\right.
$$
\n
$$
\leq -2\left(1-\frac{16n^{-0.5}\left\|E\right\|}{\Delta}\right)(\hat{u}_{1,-i}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*}))\right)
$$
\n
$$
\geq \mathbb{I}\left\{\left(1+\frac{C_{1}\beta^{-0.5}n^{-0.5}\left\|E\right\|}{\Delta}\right)\Delta\leq -2(\hat{u}_{1,-i}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*}))\right\}.
$$

Since $\theta_a^* = -\theta_{z_i^*}^*$, we have $sign(u_1^T(\theta_{z_i^*}^* - \theta_a^*)) = sign(u_1^T\theta_{z_i^*}^*)$. The proof is complete. \Box

PROOF OF THEOREM [3.4.](#page-0-9) Recall that $\lambda_1 = \Delta \sqrt{n}/2$. Same as the proof of Theorem [3.1,](#page-0-22) we work on the high-probability event (46) .

For the upper bound, from Lemma [3.2,](#page-0-11) there exists some $\phi \in \Phi$ such that for any $i \in [n]$, $\mathbb{I}\left\{\hat{z}_i \neq \phi(z_i^*)\right\} \leq \mathbb{I}\left\{\left(1 - C_1 \psi_3^{-1}\right) \Delta \leq 2\left\|\hat{u}_{1, -i}\hat{u}_{-i}^T \epsilon_i\right\|\right\} = \mathbb{I}\left\{\left(1 - C_1 \psi_3^{-1}\right) \Delta \leq 2\left|\hat{u}_{1, -i}^T \epsilon_i\right|\right\},\$ for some $C_1 > 0$, where the last inequality is due to that ψ_3 is large. By Davis-Kahan Theorem, we know there exists some $s_i \in \{-1, 1\}$ such that $\|\hat{u}_{1,-i} - s_i u_1\| \leq 2 \|E\|/(\sqrt{n-1}\Delta) \leq$ $4\psi_3^{-1}$. Since $\langle \hat{u}_{1,-i}, u_1 \rangle \ge 0$ is assumed, we have $s_i = 1$ for all $i \in [n]$. Then

$$
\mathbb{I}\left\{\hat{z}_{i} \neq \phi(z_{i}^{*})\right\} \leq \mathbb{I}\left\{\left(1 - C_{1}\psi_{3}^{-1}\right)\Delta \leq 2\left|u_{1}^{T}\epsilon_{i}\right| + 2\left|(\hat{u}_{1, -i} - s_{i}u_{1})^{T}\epsilon_{i}\right|\right\}
$$
\n
$$
\leq \mathbb{I}\left\{\left(1 - (C_{1} + C_{2})\psi_{3}^{-1}\right)\Delta \leq 2\left|u_{1}^{T}\epsilon_{i}\right|\right\} + \mathbb{I}\left\{C_{2}\psi_{3}^{-1}\Delta \leq 2\left|(\hat{u}_{1, -i} - s_{i}u_{1})^{T}\epsilon_{i}\right|\right\},
$$

where $C_2 > 0$ is a constant whose value will be determined later. Due to the independence of $\hat{u}_{1,-i} - s_i u_1$ and ϵ_i , we have $(\hat{u}_{1,-i} - s_i u_1)^T \epsilon_i \sim SG(16\psi_3^{-2}\sigma^2)$ and then

$$
\mathbb{EI}\left\{C_2\Delta \leq 2\left|\left(\hat{u}_{1,-i} - s_i u_1\right)^T \epsilon_i\right|\right\} \leq 2\exp\left(-\frac{C_2^2\Delta^2}{128\sigma^2}\right).
$$

On the other hand, $u_1^T \epsilon_i = p^{-\frac{1}{2}} \sum_{j=1}^p \epsilon_{i,j}$ where $\{\epsilon_{i,j}\}_{j \in [p]}$ are i.i.d. with variance $\bar{\sigma}^2$, which can be approximated by a normal distribution. Since the distribution F is sub-Gaussian, its moment generating function exists. Then we can use the following KMT quantile inequality (see Proposition [KMT] of [\[30\]](#page-0-43)). Let $Y \stackrel{d}{=} \bar{\sigma}^{-1} p^{-\frac{1}{2}} \sum_{j=1}^p \epsilon_{i,j}$. There exist some constants $D, \eta > 0$ and $Z \sim \mathcal{N}(0, 1)$, such that whenever $|Y| \leq \eta \sqrt{p}$, we have

$$
|Y-Z|\leq \frac{DY^2}{\sqrt{p}}+\frac{D}{\sqrt{p}}.
$$

Then,

$$
\mathbb{E}\left[\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\Delta \le 2\left|u_1^T\epsilon_i\right|\right] \n= \mathbb{E}\left[\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \le 2|Y|\right] \n\le \mathbb{E}\left[\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \le 2|Z| + \frac{2DY^2}{\sqrt{p}} + \frac{2D}{\sqrt{p}}\right] + \mathbb{E}\left[\left|Y\right| > \eta\sqrt{p}\right] \n\le \mathbb{E}\left[\left(1 - (C_1 + C_2 + C_3 + 2D)\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \le 2|Z|\right] + \mathbb{E}\left[\left\{\frac{2DY^2}{\sqrt{p}} \ge C_3\right\} + \mathbb{E}\left[\left|Y\right| > \eta\sqrt{p}\right],
$$

where
$$
C_3 > 0
$$
 is a constant. Using the fact that $Y \sim SG(1)$ with zero mean, we have
\n
$$
\mathbb{E}[\{(1 - (C_1 + C_2)\psi_3^{-1}) \Delta \le 2 | u_1^T \epsilon_i| \}
$$
\n
$$
\le 2 \exp\left(-\frac{(1 - (C_1 + C_2 + C_3 + 2D)\psi_3^{-1})^2 \Delta^2}{8\bar{\sigma}^2}\right) + 2 \exp\left(-\frac{C_3\sqrt{p}}{4D}\right) + 2 \exp\left(-\frac{\eta^2 p}{2}\right)
$$

Then we have

$$
\mathbb{E}\ell(\check{z},z^*)
$$
\n
$$
\leq \frac{1}{n}\sum_{i=1}^n \mathbb{E}\mathbb{I}\left\{\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\Delta \leq 2\left|u_1^T\epsilon_i\right|\right\} + \frac{1}{n}\sum_{i=1}^n \mathbb{E}\mathbb{I}\left\{C_2\Delta \leq 2\left|(\hat{u}_{1,-i} - s_iu_1)^T\epsilon_i\right|\right\} + e^{-0.5n}
$$
\n
$$
\leq 2\exp\left(-\frac{\left(1 - (C_1 + C_2 + C_3 + 2D)\psi_3^{-1}\right)^2\Delta^2}{8\bar{\sigma}^2}\right)
$$
\n
$$
+ 2\exp\left(-\frac{C_2^2\Delta^2}{128\sigma^2}\right) + 2\exp\left(-\frac{C_3\sqrt{p}}{4D}\right) + 2\exp\left(-\frac{\eta^2p}{2}\right) + e^{-0.5n},
$$

where $e^{-0.5n}$ is the probability that [\(46\)](#page-0-21) does not hold. Since $\sigma \leq C\bar{\sigma}$, when C_2 is chosen to satisfy $C_2^2/(128C^2) \ge 16$, we have

$$
\mathbb{E}\ell(\check{z},z^*) \le 2\exp\left(-\frac{\left(1 - C''\psi_3^{-1}\right)^2 \Delta^2}{8\bar{\sigma}^2}\right) + \exp\left(-C''\sqrt{p}\right) + e^{-0.5n},\,
$$

for some constant $C'' > 0$.

For the lower bound, from [\(70\)](#page-0-9) we know

$$
\mathbb{I}\left\{\check{z}_i \neq \phi(z_i^*)\right\} \geq \mathbb{I}\left\{\left(1 + C_4\psi_3^{-1}\right)\Delta \leq -2(\hat{u}_{1,-i}^T\epsilon_i)\text{sign}(u_1^T(\theta_{\phi(z_i^*)}-\theta_{3-\phi(z_i^*)}))\right\},\
$$

for some constant $C_4 > 0$ assuming ψ_3 is large. Using the same argument as in the upper bound, we are going to decompose $\hat{u}_{1,-i}^T \epsilon_i$ into $u_1^T \epsilon_i$ and $(\hat{u}_{1,-i} - y_1)^T \epsilon_i$. Hence,

$$
\mathbb{I}\left\{\tilde{z}_{i} \neq \phi(z_{i}^{*})\right\} \geq \mathbb{I}\left\{\left(1 + C_{4}\psi_{3}^{-1}\right)\Delta \leq -2(u_{1}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{\phi(z_{i}^{*})} - \theta_{3-\phi(z_{i}^{*})})) - 2\left|(\hat{u}_{1,-i} - s_{i}u_{1})^{T}\epsilon_{i}\right|\right\}
$$
\n
$$
\geq \mathbb{I}\left\{\left(1 + (C_{4} + C_{5})\psi_{3}^{-1}\right)\Delta \leq -2(u_{1}^{T}\epsilon_{i})\text{sign}(u_{1}^{T}(\theta_{\phi(z_{i}^{*})} - \theta_{3-\phi(z_{i}^{*})}))\right\}
$$
\n
$$
-\mathbb{I}\left\{C_{5}\psi_{3}^{-1}\Delta \leq 2\left|(\hat{u}_{1,-i} - s_{i}u_{1})^{T}\epsilon_{i}\right|\right\},\
$$

for some constant $C_5 > 0$ whose value to be chosen. Let

$$
Y' \stackrel{d}{=} \bar{\sigma}^{-1}(u_1^T \epsilon_i) \text{sign}(u_1^T (\theta_{\phi(z_i^*)} - \theta_{3-\phi(z_i^*)})) = \text{sign}(u_1^T (\theta_{\phi(z_i^*)} - \theta_{3-\phi(z_i^*)})) \bar{\sigma}^{-1} p^{-\frac{1}{2}} \sum_{j=1}^p \epsilon_{i,j}.
$$

Then using the same argument above, there exists some $Z' \sim \mathcal{N}(0, 1)$ such that whenever $Y' \le \eta' \sqrt{p}$, we have $|Y' - Z'| \le \frac{D'Y'^2}{\sqrt{p}} + \frac{D'}{\sqrt{p}}$ where $D', \eta' > 0$ are constants. Then

$$
\mathbb{E}\mathbb{I}\left\{\left(1+(C_4+C_5)\,\psi_3^{-1}\right)\Delta\leq-2(u_1^T\epsilon_i)\text{sign}(u_1^T(\theta_{\phi(z_i^*)}-\theta_{3-\phi(z_i^*)}))\right\}
$$
\n
$$
=\mathbb{E}\mathbb{I}\left\{\left(1+(C_4+C_5)\,\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}}\leq-2Y'\right\}
$$
\n
$$
\geq \mathbb{E}\mathbb{I}\left\{\left(1+(C_4+C_5)\,\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}}\leq-2Z'-\frac{2DY'^2}{\sqrt{p}}-\frac{2d}{\sqrt{p}}\right\}\mathbb{I}\left\{Y'\leq\eta'\sqrt{p}\right\}
$$
\n
$$
\geq \mathbb{E}\mathbb{I}\left\{\left(1+(C_4+C_5+2D+C_6)\,\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}}\leq-2Z'\right\}-\mathbb{E}\mathbb{I}\left\{\frac{2DY'^2}{\sqrt{p}}\geq C_6\right\}-\mathbb{E}\mathbb{I}\left\{Y'>\eta'\sqrt{p}\right\},\right\}
$$

.

where $C_6 > 0$ is a constant. Then following the proof of the upper bound, and by a proper choice of *C*5, we have

$$
\mathbb{E}\ell(\check{z},z^*) \ge 2\exp\left(-\frac{\left(1+C'''\psi_3^{-1}\right)^2\Delta^2}{8\bar{\sigma}^2}\right) - \exp\left(-C'''\sqrt{p}\right) - e^{-0.5n},\right)
$$

for some constant $C''' > 0$.

D.2. Proofs of Lemma [3.4](#page-0-44) and Theorem [3.5.](#page-0-45)

PROOF OF LEMMA [3.4.](#page-0-44) For the upper bound, we consider the following likelihood ratio test. For any $x \in \mathbb{R}^p$, define the two log-likelihood functions as

$$
l_1(x) = \sum_{j=1}^p \log f(x_j - \delta)
$$
, and $l_2(x) = \sum_{j=1}^p \log f(x_j + \delta)$.

Then for each $i \in [n]$, define the likelihood ratio test as

$$
\hat{z}_i^{\text{LRT}} = \begin{cases} 1, & \text{if } l_1(X_i) \ge l_2(X_i), \\ 2, & \text{otherwise.} \end{cases}
$$

Then for any $i \in [n]$ such that $z_i^* = 1$, we have

$$
\mathbb{E} \left\{ \hat{z}_i^{\text{LRT}} = 2 \right\} = \mathbb{P} \left(l_2(X_i) > l_1(X_i) \right) = \mathbb{P} \left(\sum_{j=1}^p \log \frac{f(2\delta + \epsilon_{i,j})}{f(\epsilon_{i,j})} > 0 \right) = \mathbb{P} \left(\sum_{j=1}^p \log \frac{f_{\frac{\Delta}{\sqrt{p}}}(\epsilon_{i,j})}{f_0(\epsilon_{i,j})} > 0 \right),
$$

where we use the fact $2\delta = \frac{\Delta}{\sqrt{2}}$ $\frac{1}{p}$. Since Δ is a constant, by local asymptotic normality (c.f., Chapter 7, [\[41\]](#page-0-46)), we have

$$
\sum_{j=1}^p \log \frac{f_{\frac{\Delta}{\sqrt{p}}}(\epsilon_{i,j})}{f_0(\epsilon_{i,j})} \xrightarrow{d} \mathcal{N}\left(-\frac{\mathcal{I}\Delta^2}{2}, \mathcal{I}\Delta^2\right).
$$

Then, $\lim_{p\to\infty} \mathbb{E} \left[\left\{ \hat{z}_i^{\text{LRT}} = 2 \right\} \leq C_1 \exp \left(-\mathcal{I} \Delta^2 / 8 \right)$ for some constant $C_1 > 0$. We have the same upper bound if $z_i^* = 2$ instead. Hence,

$$
\lim_{p \to \infty} \inf_z \sup_{z^* \in [2]^n} \mathbb{E}\ell(z, z^*) \le \lim_{p \to \infty} \sup_{z^* \in [2]^n} \mathbb{E}\ell(\hat{z}^{\text{LRT}}, z^*) \le \exp\left(-\frac{\mathcal{I}\Delta^2}{8}\right).
$$

For the lower bound, instead of allowing $z^* \in [2]^n$, we consider a slightly smaller parameter space. Define $\mathcal{Z} = \{z \in [2]^n : z_i = 1, \forall 1 \le i \le n/3, z_i = 2, \forall n/3 + 1 \le i \le 2n/3\}.$ Then for any $z, z' \in \mathcal{Z}$ we have $\ell(z, z') = n^{-1} \sum_{i=1}^n \mathbb{I} \{z_i \neq z'_i\} \leq 1/3$ due to the fact $n^{-1} \sum_{i=1}^{n} \mathbb{I} \{ \phi(z_i) \neq z'_i \} \geq 1/3$ if $\phi \neq$ Id. Hence,

$$
\inf_{z} \sup_{z^* \in [2]^n} \mathbb{E}\ell(z, z^*) \ge \inf_{z} \sup_{z^* \in \mathcal{Z}} \mathbb{E}\ell(z, z^*) \ge n^{-1} \inf_{z} \sup_{z^* \in \mathcal{Z}} \mathbb{E} \sum_{i \in [n]} \mathbb{I} \{z_i \neq z_i^* \}
$$

$$
\ge n^{-1} \sum_{i > 2n/3} \inf_{z_i} \sup_{z_i^* \in [2]} \mathbb{E} \{z_i \neq z_i^* \} = \frac{1}{3} \inf_{z_n} \sup_{z_n^* \in [2]} \mathbb{E} \{z_n \neq z_n^* \},
$$

where it is reduced into a testing problem on whether X_n has mean θ_1^* or θ_2^* . According to the Neyman-Pearson Lemma, the optimal procedure is the likelihood ratio test \hat{z}_n^{LRT} defined above. By the same argument, we have

$$
\liminf_{p \to z} \sup_{z^* \in [2]^n} \mathbb{E}\ell(z, z^*) \ge \frac{1}{3} \liminf_{p \to z_n} \sup_{z_n^* \in [2]} \mathbb{E}\{z_n \ne z_n^*\} \ge C_2 \exp\left(-\frac{\mathcal{I}\Delta^2}{8}\right),
$$

for some constant $C_2 > 0$.

 \Box

PROOF OF THEOREM [3.5.](#page-0-45) First, we have the following connection between the Fisher information $\mathcal I$ and the variance $\bar{\sigma}^2$:

$$
\mathcal{I}\bar{\sigma}^2 = \left(\int \left(\frac{f'}{f}\right)^2 f \mathrm{d}x\right) \left(\int x^2 f \mathrm{d}x\right) \ge \left(\int \frac{f'}{f} x f \mathrm{d}x\right)^2 = \left(\int x f' \mathrm{d}x\right)^2 = 1,
$$

where we use Cauchy-Schwarz inequality and the integral by part $\int x f' dx = \int x f dx$ $\int f dx = 0 - 1 = -1$. The equation holds if and only if $f'/f \propto x$, which is equivalent to *F* being normally distributed.

APPENDIX E: AUXILIARY LEMMAS AND PROPOSITIONS AND THEIR PROOFS

PROPOSITION E.1. *For Y* and \hat{Y} defined in [\(1\)](#page-0-47), we have [\(2\)](#page-0-48) holds assuming $\sigma_r - \sigma_{r+1}$ > $2 || (I - U_r U_r^T) y_n ||.$

PROOF. Recall the augmented matrix Y' is defined as $(Y, U_r U_r^T y_n)$. Note that $U_r U_r^T Y$ is the best rank-*r* approximation of *Y* . Since

$$
\left\| \left(I - U_r U_r^T \right) Y' \right\|_{\rm F} = \left\| \left(\left(I - U_r U_r^T \right) Y, 0 \right) \right\|_{\rm F} = \left\| \left(I - U_r U_r^T \right) Y \right\|_{\rm F},
$$

we have $U_r U_r^T Y'$ also being the best rank-*r* approximation of Y'. This proves that span(U_r) and $U_r U_r^T$ are also the leading r left singular subspace and projection matrix of Y'. Then $\hat{U}_r \hat{U}_r^T - U_r U_r^T$ is about the perturbation between \hat{Y} and Y' .

Let σ'_r , σ'_{r+1} be the *r*th and $(r+1)$ th largest singular values of Y' , respectively. By Wedin's Thereom (see Section 2.3 of [\[9\]](#page-0-49)), if $\sigma'_r - \hat{\sigma}_{r+1} > 0$, then we have

(73)
$$
\|\sin \Theta(\hat{U}_r, U_r)\|_{\rm F} \leq \frac{\left\|\hat{Y} - Y'\right\|_{\rm F}}{\sigma'_r - \hat{\sigma}_{r+1}} = \frac{\left\|\left(I - U_r U_r^T\right)y_n\right\|}{\sigma'_r - \hat{\sigma}_{r+1}}.
$$

Regarding the values of σ'_r and σ'_{r+1} , first we have $\sigma'_r \ge \sigma_r$. This is because

$$
\sigma'_r = \inf_{x \in \text{span}(U_r)} ||x^T Y'|| = \inf_{x \in \text{span}(U_r)} ||(x^T Y, x^T y_n)|| \ge \inf_{x \in \text{span}(U_r)} ||x^T Y|| \ge \sigma_r.
$$

In addition, we have $\sigma'_{r+1} = \sigma_{r+1}$, due to the fact that $(I - U_r U_r^T) Y' = ((I - U_r U_r^T) Y, 0)$. By Weyl's inequality, we have

$$
|\hat{\sigma}_{r+1} - \sigma'_{r+1}| \le ||Y - Y'|| = ||(I - U_r U_r^T) y_n||.
$$

Hence, if $\sigma_r - \sigma_{r+1} > 2 ||(I - U_r U_r^T) y_n||$ is further assumed, we have

(74)
$$
\sigma'_{r} - \hat{\sigma}_{r+1} \geq \sigma_{r} - \sigma_{r+1} - ||(I - U_{r}U_{r}^{T})y_{n}|| \geq \frac{1}{2} (\sigma_{r} - \sigma_{r+1}).
$$

With [\(73\)](#page-0-50), [\(74\)](#page-0-16), and the fact $\|\hat{U}_r\hat{U}_r^T - U_rU_r^T\|_F = \sqrt{2}\|\sin\Theta(\hat{U}_r, U_r)\|_F$ (see Lemma 1 of [\[9\]](#page-0-49)), the proof is complete.

LEMMA E.1. Let $E = (\epsilon_1, \ldots, \epsilon_n) \in \mathbb{R}^{p \times n}$ *be a random matrix with each column* $\epsilon_i \sim$ $SG_p(\sigma^2)$, $\forall i \in [n]$ *independently. Then*

$$
\mathbb{P}\left(\|E\| \ge 4t\sigma(\sqrt{n} + \sqrt{p})\right) \le \exp\left(-\frac{(t^2-3)n}{2}\right),\,
$$

for any $t > 2$ *.*

PROOF. We follow a standard ϵ -net argument. Let *U* and *V* be a 1/4 covering set of the unit sphere in \mathbb{R}^p and in \mathbb{R}^n , respectively. That is, for any $u \in \mathbb{R}^p$ such that $||u|| = 1$, there exists a $u' \in \mathcal{U}$ such that $||u'|| = 1$ and $||u - u'|| \le 1/4$. Similarly, for any $v \in \mathbb{R}^n$ such that $||v|| = 1$, there exists a $v' \in V$ such that $||v'|| = 1$ and $||v - v'|| \le 1/4$. Then

$$
|u^T E v| = |u^T E v' + u^T E (v - v') + (u - u')^T E v' + (u - u')^T E (v - v')|
$$

$$
\le |u^T E v'| + |u^T E (v - v')| + |(u - u')^T E v'| + |(u - u')^T E (v - v')|.
$$

Maximizing over *u, v* on both sides, we have

$$
||E|| = \max_{u \in \mathbb{R}^p, v \in \mathbb{R}^n : ||u|| = ||v|| = 1} |u^T E v| \le \max_{u' \in \mathcal{U}, v' \in \mathcal{V}} \left| u'^T E v' \right| + \frac{1}{4} ||E|| + \frac{1}{4} ||E|| + \frac{1}{16} ||E||.
$$

Hence,

$$
||E||\leq 4 \max_{u'\in\mathcal{U},v'\in\mathcal{V}}\left|u^{'T}Ev'\right|.
$$

For any $u' \in \mathcal{U}$, $v' \in \mathcal{V}$, we have each $u'^T \epsilon_i$ being an independent $SG(\sigma^2)$ and then $u'^T Ev' \sim$ $SG(\sigma^2)$. Note $|U| \leq 9^p \leq e^{3p}$ and similarly $|V| \leq e^{3n}$. Then by the tail probability of sub-Gaussian random variable and by the union bound, we have

$$
\mathbb{P}\left(\|E\| \le 4t\sigma(\sqrt{n} + \sqrt{p})\right) \le \mathbb{P}\left(\max_{u' \in \mathcal{U}, v' \in \mathcal{V}} \left|u'^T E v'\right| \le t\sigma(\sqrt{n} + \sqrt{p})\right)
$$

$$
\le |U| \, |V| \exp\left(-\frac{t^2\left(\sqrt{n} + \sqrt{p}\right)^2}{2}\right)
$$

$$
\le \exp\left(-\frac{(t^2 - 3)n}{2}\right),
$$

for any $t \geq 2$.

LEMMA E.2. Let $X \sim SG_d(\sigma^2)$. Consider any $k \leq d$. For any matrix $U = (u_1, \ldots, u_k) \in$ $\mathbb{R}^{d \times k}$ *that is independent of* \overline{X} *and is with orthogonal columns* $\{u_i\}_{i \in [k]}$ *. We have*

 \Box

$$
\mathbb{P}\left(\left\|UU^TX\right\|^2 \ge \sigma^2(k + 2\sqrt{kt} + 2t)\right) \le e^{-t}.
$$

PROOF. Note that $tr(UU^T) = tr((UU^T)^2) = k$ and $||UU^T|| = 1$. This is a direct consequence of Theorem 1 in [\[18\]](#page-0-51) for concentration of quadratic forms of sub-Gaussian random vectors. \Box

PROOF OF PROPOSITION [3.1.](#page-0-9) Define $\hat{P} = \sum_{i \in [r]} \hat{\lambda}_i \hat{u}_i \hat{v}_i^T$. Due to the fact that \hat{P} is the best rank-r approximation of X in spectral norm and P is rank- κ , under the assumption that $\kappa \leq r$, we have that

$$
\left\| \hat{P} - X \right\| \le \|P - X\| = \|E\|.
$$

Since $r \leq k$ is assumed, the rank of $\hat{P} - P$ his at most 2*k*, and we have

$$
(75) \qquad \left\| \hat{P} - P \right\|_{\mathrm{F}} \le \sqrt{2k} \left\| \hat{P} - P \right\| \le \sqrt{2k} \left(\left\| \hat{P} - X \right\| + \left\| P - X \right\| \right) \le 2\sqrt{2k} \left\| E \right\|
$$

Now, denote $\hat{\Theta} := (\hat{\theta}_{\hat{z}_1}, \hat{\theta}_{\hat{z}_2}, \dots, \hat{\theta}_{\hat{z}_n})$. Since $\hat{\Theta}$ is the solution to the *k*-means objective [\(15\)](#page-0-52), we have that

$$
\left\| \hat{\Theta} - \hat{P} \right\|_{\rm F} \le \left\| P - \hat{P} \right\|_{\rm F}
$$

.

Hence, by the triangle inequality, we obtain that

$$
\left\|\hat{\Theta} - P\right\|_{\mathrm{F}} \leq 2\left\|\hat{P} - P\right\|_{\mathrm{F}} \leq 4\sqrt{2k} \left\|E\right\|.
$$

Now, define the set *S* as

$$
S = \left\{ i \in [n] : \left\| \hat{\theta}_{\hat{z}_i} - \theta^*_{z_i^*} \right\| > \frac{\Delta}{2} \right\}.
$$

Since $\left\{\hat{\theta}_{\hat{z}_i} - \theta_{z_i^*}^*\right\}$ \mathcal{L} are exactly the columns of $\hat{\Theta} - P$, we have that

$$
|S| \le \frac{\left\|\hat{\Theta} - P\right\|_{\mathrm{F}}^2}{\left(\Delta/2\right)^2} \le \frac{128k\left\|E\right\|^2}{\Delta^2}.
$$

Under the assumption (16) we have

$$
\frac{\beta \Delta^2 n}{k^2 \left\| E \right\|^2} \ge 256,
$$

which implies

$$
|S| \le \frac{\beta n}{2k}.
$$

We now show that all the data points in S^C are correctly clustered. We define

$$
C_j = \left\{ i \in [n] : z_i^* = j, i \in S^C \right\}, \ j \in [k].
$$

The following holds:

- For each $j \in [k]$, C_j cannot be empty, as $|C_j| \ge |\{i : z_i^* = j\}| |S| \ge 0$.
- For each pair $j, l \in [k], j \neq l$, there cannot exist some $i \in C_j, i' \in C_l$ such that $\hat{z}_i = \hat{z}_{i'}$. Otherwise $\hat{\theta}_{\hat{z}_i} = \hat{\theta}_{\hat{z}_i}$, which would imply

$$
\begin{aligned} \left\|\theta_j^* - \theta_l^*\right\| &= \left\|\theta_{z_i^*}^* - \theta_{z_{i'}^*}^*\right\| \\ &\leq \left\|\theta_{z_i^*}^* - \hat{\theta}_{\hat{z}_i}\right\| + \left\|\hat{\theta}_{\hat{z}_i} - \hat{\theta}_{\hat{z}_{i'}}\right\| + \left\|\hat{\theta}_{\hat{z}_{i'}} - \theta_{z_{i'}^*}^*\right\| < \Delta, \end{aligned}
$$

contradicting with the definition of Δ .

Since \hat{z}_i can only take values in [k], we conclude that the sets $\{\hat{z}_i : i \in C_j\}$ are disjoint for all $j \in [k]$. That is, there exists a permutation $\phi \in \Phi$, such that

$$
\hat{z}_i = \phi(j), \ i \in C_j, \ j \in [k].
$$

This implies that $\sum_{i \in S^C} \mathbb{I} \{ \hat{z}_i \neq \phi(z_i^*) \} = 0$. Hence, we obtain that

$$
|\{i \in [n]: \hat{z}_i \neq \phi(z_i^*)\}| \leq |S| \leq \frac{128k ||E||^2}{\Delta^2}.
$$

Since $|S| \leq \frac{\beta n}{2k}$ (which means $\ell(\hat{z},z^*) \leq \frac{\beta n}{2k}$ from the above display), for any $\psi \in \Phi$ such that $\psi \neq \phi$, we have $|\{i \in [n]: \hat{z}_i \neq \psi(z_i^*)\}| \geq 2\beta n/k - |S| \geq \beta n/k$. As a result, we have

$$
\ell(\hat{z}, z^*) = \frac{1}{n} |\{i \in [n] : \hat{z}_i \neq \phi(z_i^*)\}| \le \frac{128k \|E\|^2}{n\Delta^2}.
$$

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Moreover, for each $a \in [k],$ we have

$$
\left\|\hat{\theta}_{\phi(a)} - \theta_a^*\right\|^2 \le \frac{\left\|\hat{\Theta} - P\right\|_{\mathrm{F}}^2}{|\{i \in [n] : \hat{z}_i = \phi(a), z_i^* = a\}|} \le \frac{\left\|\hat{\Theta} - P\right\|_{\mathrm{F}}^2}{\frac{\beta n}{k} - |S|} \le \frac{64k^2 \|E\|^2}{\beta n}
$$