# SUPPLEMENT TO "LEAVE-ONE-OUT SINGULAR SUBSPACE PERTURBATION ANALYSIS FOR SPECTRAL CLUSTERING"

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#### APPENDIX A: PROOF OF THEOREM 2.3

The proof idea is similar to that of Theorem 2.2 but with more involved calculation as r is not necessarily  $\kappa$ . Consider any  $i \in [n]$ . Define

$$\tilde{\rho}_{-i} := \frac{\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1}}{\left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) X_i \right\|}$$

We need to verify  $\tilde{\rho}_{-i} > 2$  first in order to apply Theorem 2.1. Recall the definition of  $P_{-i}$  in (36) and  $E_{-i}$  in (38). Let the SVD of  $P_{-i}$  be

$$P_{-i} = \sum_{j=1}^{p \wedge (n-1)} \lambda_{-i,j} u_{-i,j} v_{-i,j}^T,$$

where  $\lambda_{-i,1} \ge \lambda_{-i,2} \ge \ldots \ge \lambda_{-i,p \land (n-1)}$ . Denote  $U_{-i,1:r} = (u_{-i,1}, u_{-i,2}, \ldots, u_{-i,r}) \in \mathbb{O}^{p \times r}$ . Then by Weyl's inequality, we have

(48) 
$$|\hat{\lambda}_{-i,r} - \lambda_{-i,r}|, |\hat{\lambda}_{-i,r+1} - \lambda_{-i,r+1}| \le ||E_{-i}|| \le ||E||.$$

Then the numerator

(49) 
$$\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1} \ge \lambda_{-i,r} - \lambda_{-i,r+1} - 2 \|E\|.$$

In the following, we are going to connect  $\lambda_{-i,r} - \lambda_{-i,r+1}$  with  $\lambda_r - \lambda_{r+1}$ . To bridge the gap between  $\lambda_{-i,r}$ ,  $\lambda_{-i,r+1}$  and  $\lambda_r$ ,  $\lambda_{r+1}$ , define

$$\tilde{P}_{-i} := (\theta_{z_1^*}^*, \dots, \theta_{z_{i-1}^*}^*, U_{-i,1:r}U_{-i,1:r}^T \theta_{z_i^*}^*, \theta_{z_{i+1}^*}^*, \dots, \theta_{z_n^*}^*) \in \mathbb{R}^{p \times n}$$

Let  $\tilde{\lambda}_{-i,1} \geq \tilde{\lambda}_{-i,2} \geq \ldots \geq \tilde{\lambda}_{-i,p \wedge n}$  be its singular values. Note that  $U_{-i,1:r}U_{-i,1:r}^T \tilde{P}_{-i}$  is the best rank-*r* approximation of  $\tilde{P}_{-i}$ . This is because for any rank-*r* projection matrix  $M \in \mathbb{R}^{p \times p}$  such that  $M^2 = M$ , we have

$$\begin{split} \left\| \tilde{P}_{-i} - MM^{T} \tilde{P}_{-i} \right\|_{\mathrm{F}}^{2} &= \left\| (I - MM^{T}) P_{-i} \right\|_{\mathrm{F}}^{2} + \left\| (I - MM^{T}) U_{-i,1:r} U_{-i,1:r}^{T} \theta_{z_{i}^{*}}^{*} \right\|_{\mathrm{F}}^{2} \\ &\geq \left\| (I - U_{-i,1:r} U_{-i,1:r}^{T}) P_{-i} \right\|_{\mathrm{F}}^{2} + 0 \\ &= \left\| \tilde{P}_{-i} - U_{-i,1:r} U_{-i,1:r}^{T} \tilde{P}_{-i} \right\|_{\mathrm{F}}^{2}, \end{split}$$

where we use the fact  $U_{-i,1:r}U_{-i,1:r}^TP_{-i}$  is the best rank-*r* approximation of  $P_{-i}$ . Hence, span $(U_{-i,1:r})$  is exactly the leading *r* left singular space of  $\tilde{P}_{-i}$ . It immediately implies:

•  $\tilde{\lambda}_{-i,j} = \lambda_{-i,j}$  for any  $j \ge r+1$ , including

$$\tilde{\lambda}_{-i,r+1} = \lambda_{-i,r+1}$$

• Since  $U_{-i,1:r}U_{-i,1:r}^T \tilde{P}_{-i}$  and  $U_{-i,1:r}U_{-i,1:r}^T P_{-i}$  only differ by one column where the latter one can be seen as the leave-one-out counterpart of the former one, using the same argument as in (37), we have

(51) 
$$\lambda_{-i,r}^2 \ge \left(1 - \frac{k}{\beta n}\right) \tilde{\lambda}_{-i,r}^2.$$

Then from (49), we have

(52) 
$$\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1} \ge \sqrt{1 - \frac{k}{\beta n}} \tilde{\lambda}_{-i,r} - \tilde{\lambda}_{-i,r+1} - 2 \|E\|.$$

For the difference between  $\tilde{\lambda}_{-i,r}$ ,  $\tilde{\lambda}_{-i,r+1}$  and  $\lambda_r$ ,  $\lambda_{r+1}$ , we use the Weyl's inequality again:

$$\max_{j \in [k]} \left| \tilde{\lambda}_{-i,j} - \lambda_j \right| \le \left\| P - \tilde{P}_{-i} \right\| = \left\| \theta_{z_i^*}^* - U_{-i,1:r} U_{-i,1:r}^T \theta_{z_i^*}^* \right\|$$

In the proof of Theorem 2.2, we show  $u_{-i,j} \in \text{span}(\{\theta_a^*\}_{a \in [k]})$  for each  $j \in [\kappa]$ . Then

$$\begin{split} \left\| \theta_{z_{i}^{*}}^{*} - U_{-i,1:r} U_{-i,1:r}^{T} \theta_{z_{i}^{*}}^{*} \right\| &= \left\| \left( u_{-i,r+1}, \dots, u_{-i,\kappa} \right) \left( u_{-i,r+1}, \dots, u_{-i,\kappa} \right)^{T} \theta_{z_{i}^{*}}^{*} \right\| \\ &= \sqrt{\sum_{a \in [\kappa]: a \ge r+1} \left( u_{-i,a}^{T} \theta_{z_{i}^{*}}^{*} \right)^{2}}. \end{split}$$

For any  $a \in [\kappa]$  such  $a \ge r+1$ , we have

$$\begin{split} \left(u_{-i,a}^{T}\theta_{z_{i}^{*}}^{*}\right)^{2} &\leq \frac{1}{\left|\left\{j\in[n]:z_{j}^{*}=z_{i}^{*}\right\}\right|-1}\sum_{j\in[n]:j\neq i, z_{j}^{*}=z_{i}^{*}}\left(u_{-i,a}^{T}\theta_{z_{j}^{*}}^{*}\right)^{2} \leq \frac{1}{\frac{\beta n}{k}-1}(u_{-i,a}^{T}P_{-i})^{2}\\ &\leq \frac{\lambda_{-i,a}^{2}}{\frac{\beta n}{k}-1} \leq \frac{\lambda_{-i,r+1}^{2}}{\frac{\beta n}{k}-1}. \end{split}$$

Hence, we obtain  $\|\theta_{z_i^*}^* - U_{-i,1:r}U_{-i,1:r}^T\theta_{z_i^*}^*\| \leq \sqrt{\kappa}\lambda_{-i,a}/\sqrt{\beta n/k - 1}$  and consequently,

(53) 
$$\max_{j\in[k]} \left| \tilde{\lambda}_{-i,j} - \lambda_j \right| \le \frac{\sqrt{\kappa\lambda_{-i,r+1}}}{\sqrt{\frac{\beta n}{k} - 1}}.$$

Then together with (50), we have  $|\lambda_{-i,r+1} - \lambda_{r+1}| \le \sqrt{\kappa}\lambda_{-i,r+1}/\sqrt{\beta n/k - 1}$  and hence

(54) 
$$\lambda_{-i,r+1} \le \frac{\lambda_{r+1}}{1 - \frac{\sqrt{\kappa}}{\sqrt{\frac{\beta_n}{k} - 1}}}.$$

Denote  $d := \beta n/k$ . With (52), we have

$$\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1} \ge \sqrt{\frac{d-1}{d}} \left( \lambda_r - \frac{\lambda_{-i,r+1}}{\sqrt{d-1}} \right) - \left( \lambda_{r+1} + \frac{\lambda_{-i,r+1}}{\sqrt{d-1}} \right) - 2 \|E\|$$

$$\ge \sqrt{\frac{d-1}{d}} \lambda_r - \lambda_{r+1} \left( 1 + \left( \frac{1}{\sqrt{d}} + \frac{1}{\sqrt{d-1}} \right) \frac{1}{1 - \frac{\sqrt{\kappa}}{\sqrt{d-1}}} \right) - 2 \|E\|$$

$$\ge \sqrt{\frac{d-1}{d}} \left( \lambda_r - \lambda_{r+1} - \frac{4}{\sqrt{d}} \lambda_{r+1} \right) - 2 \|E\|$$

$$\ge \frac{3}{4} \left( \lambda_r - \lambda_{r+1} - \frac{4}{\sqrt{d}} \lambda_{r+1} \right) - 2 \|E\|,$$
(55)

where in the last two inequalities we use the assumption that  $d/k \ge 10$ . As a consequence, we have

$$\tilde{\rho}_{-i} \ge \frac{\hat{\lambda}_{-i,r} - \hat{\lambda}_{-i,r+1}}{\left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) X_i \right\|} \ge \frac{\frac{3}{4} \left( \lambda_r - \lambda_{r+1} - \frac{4}{\sqrt{d}} \lambda_{r+1} \right) - 2 \left\| E \right\|}{\left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) X_i \right\|}.$$

Next, we are going to simplify the denominator of the above display. Using the orthogonality of the singular vectors, we have

$$\begin{split} & \left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^{T} \right) \theta_{z_{i}^{*}}^{*} \right\| \\ & \leq \left\| \left( I - \hat{U}_{-i,1:\kappa} \hat{U}_{-i,1:\kappa}^{T} \right) \theta_{z_{i}^{*}}^{*} \right\| + \left\| (\hat{u}_{-i,r+1}, \dots, \hat{u}_{-i,\kappa}) \left( \hat{u}_{-i,r+1}, \dots, \hat{u}_{-i,\kappa} \right)^{T} \theta_{z_{i}^{*}}^{*} \right\| \\ & = \left\| \left( I - \hat{U}_{-i,1:\kappa} \hat{U}_{-i,1:\kappa}^{T} \right) \theta_{z_{i}^{*}}^{*} \right\| + \sqrt{\sum_{j=r+1}^{\kappa} \left( \hat{u}_{-i,j}^{T} \theta_{z_{i}^{*}}^{*} \right)^{2}} \\ & \leq \frac{3\sqrt{\kappa} \|E\|}{\sqrt{\frac{\beta n}{k} - 1}} + \sqrt{\sum_{j=r+1}^{\kappa} \left( \frac{\hat{\lambda}_{-i,j}}{\sqrt{\frac{\beta n}{k} - 1}} + \frac{\|E\|}{\sqrt{\frac{\beta n}{k} - 1}} \right)^{2}} \\ & \leq \frac{3\sqrt{\kappa} \|E\|}{\sqrt{\frac{\beta n}{k} - 1}} + \sqrt{\kappa} \left( \frac{\hat{\lambda}_{-i,r+1}}{\sqrt{\frac{\beta n}{k} - 1}} + \frac{\|E\|}{\sqrt{\frac{\beta n}{k} - 1}} \right), \end{split}$$

where the second to the inequality is due to (41) and (44). By (54) and the Weyl's inequality, we have

$$\hat{\lambda}_{-i,r+1} \le \lambda_{-i,r+1} + \|E\| \le \frac{1}{1 - \frac{\sqrt{\kappa}}{\sqrt{\frac{\beta n}{k} - 1}}} \lambda_{r+1} + \|E\|.$$

Then, with the assumption  $\beta n/k^2 \ge 10$ , we have

$$\begin{split} \left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) \theta_{z_i^*}^* \right\| &\leq \frac{3\sqrt{\kappa} \|E\|}{\sqrt{\frac{\beta n}{k} - 1}} + \sqrt{\kappa} \left( \frac{\lambda_{r+1}}{\sqrt{\frac{\beta n}{k} - 1} - \sqrt{\kappa}} + \frac{2 \|E\|}{\sqrt{\frac{\beta n}{k} - 1}} \right) \\ &\leq \frac{\sqrt{k\kappa}}{\sqrt{\beta n}} (6 \|E\| + 2\lambda_{r+1}). \end{split}$$

Hence,

$$\begin{split} \left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) X_i \right\| &\leq \left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) \theta_{z_i^*}^* \right\| + \left\| \left( I - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right) \epsilon_i \right\| \\ &\leq \frac{\sqrt{k\kappa}}{\sqrt{\beta n}} (6 \|E\| + 2\lambda_{r+1}) + \|E\| \,. \end{split}$$

As a result,

$$\tilde{\rho}_{-i} \ge \frac{\frac{3}{4} \left(\lambda_r - \lambda_{r+1} - \frac{4}{\sqrt{\beta n/k}} \lambda_{r+1}\right) - 2 \left\|E\right\|}{\frac{\sqrt{k\kappa}}{\sqrt{\beta n}} (6 \left\|E\right\| + 2\lambda_{r+1}) + \left\|E\right\|} \ge \frac{\tilde{\rho}_0}{8} > 2,$$

under the assumption that  $\beta n/(k^2) \ge 10$  and (11).

The remaining part of the proof is to study  $\{\hat{u}_{-i,a}^T X_i\}_{a \in [r]}$  and then apply Theorem 2.1. Following the exact argument as in the proof of Theorem 2.2, we have

$$\sqrt{\sum_{a \in r} \left(\frac{\hat{u}_{-i,a}^T X_i}{\hat{\lambda}_{-i,a}}\right)^2} \le \frac{\sqrt{r}}{\sqrt{\frac{\beta n}{k} - 1}} + \frac{1}{\hat{\lambda}_{-i,r}} \frac{\|E\|\sqrt{r}}{\sqrt{\frac{\beta n}{k} - 1}} + \frac{1}{\hat{\lambda}_{-i,r}} \left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T \epsilon_i\right\|.$$

Under the assumption that  $\beta n/(k^2) \ge 10$  and (11), (55) is lower bounded by  $\lambda_r/2$ . This also implies  $\hat{\lambda}_{-i,r} \ge \lambda_r/2$ . Then a direct application of Theorem 2.1 leads to

$$\begin{split} \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right\|_{\mathcal{F}} &\leq \frac{4\sqrt{2}}{\tilde{\rho}_{-i}} \left( \frac{\sqrt{r}}{\sqrt{\beta n/k - 1}} + \frac{1}{\hat{\lambda}_{-i,r}} \left( \frac{\sqrt{r} \, \|E\|}{\sqrt{\beta n/k - 1}} + \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \hat{\epsilon}_i \right\| \right) \right) \\ &\leq \frac{128}{\tilde{\rho}_0} \left( \frac{\sqrt{kr}}{\sqrt{\beta n}} + \frac{\left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \hat{\epsilon}_i \right\|}{\lambda_r} \right). \end{split}$$

#### APPENDIX B: PROOFS OF RESULTS IN SECTION 3.4

Before presenting the proof of Lemma 3.3, we first show  $\hat{r}$  defined in (23) always exists. In addition, since  $\hat{r} \in [k]$  is a random variable, we are going to associate it with some deterministic set in [k]. Recall  $\lambda_1 \geq \lambda_2 \geq \ldots \geq \lambda_{p \wedge n}$  are singular values of the signal matrix P and  $\kappa$  is the its rank. Let its SVD be  $P = \sum_{i \in [p \wedge n]} \lambda_i u_i v_i^T$  with  $\{u_j\}_{j \in [p \wedge n]} \in \mathbb{R}^p$  being its left singular vectors.

LEMMA B.1. Under the same conditions as stated in Lemma 3.3,  $\hat{r}$  always exists. Furthermore, we have  $\hat{r} \in \mathcal{R}$  where

(56) 
$$\mathcal{R} := \{ a \in [k] : \lambda_a - \lambda_{a+1} \ge (\tilde{\rho} - 2) \| E \| \text{ and } \lambda_{a+1} \le (k\tilde{\rho} + 1) \| E \| \}$$

PROOF. The existence of  $\hat{r}$  can be proved by contradiction. If  $\hat{r}$  does not exist, it means that  $\{a \in [k] : \hat{\lambda}_a - \hat{\lambda}_{a+1} \ge T\}$  is empty, which implies  $\hat{\lambda}_1 < \hat{\lambda}_{k+1} + kT = \hat{\lambda}_{k+1} + k\tilde{\rho}||E||$ . By Weyl's inequality, we have  $|\hat{\lambda}_a - \lambda_a| \le ||E||$  for all singular values of X and P. Then we have  $\lambda_1 < (k\tilde{\rho} + 1)||E||$ . On the other hand, we have

$$\lambda_{1}^{2} = \max_{w \in \mathbb{R}^{p}: \|w\| = 1} \|w^{T}P\|^{2} \ge \max_{a, b \in [k]: a \neq b} \max_{w \in \mathbb{R}^{p}: \|w\| = 1} \frac{\beta n}{k} \left( \|w^{T}\theta_{a}^{*}\|^{2} + \|w^{T}\theta_{b}^{*}\|^{2} \right)$$
$$\ge \max_{a, b \in [k]: a \neq b} \max_{w \in \mathbb{R}^{p}: \|w\| = 1} \frac{\beta n}{2k} \|w^{T}\theta_{a}^{*} - w^{T}\theta_{b}^{*}\|^{2} = \frac{\beta n}{2k} \Delta^{2},$$

where the first inequality is due to the mixture model structure in P and the second inequality is due to  $2(x_1 + x_2)^2 \ge (x_1 - x_2)^2$  for any two scalars  $x_1, x_2$ . Then we have  $\lambda_1 \ge \sqrt{\beta n/(2k)} \Delta = (\tilde{\psi}_0/\sqrt{2})k^{1.5} ||E||$  by (25). Since  $\tilde{\rho} < \tilde{\psi}_0/64$  is assumed, we have  $(k\tilde{\rho}+1)||E|| < (\tilde{\psi}_0/\sqrt{2})k^{1.5} ||E||$ , which is a contradiction.

To prove the second statement, note that we have  $\hat{\lambda}_{\hat{r}} - \hat{\lambda}_{\hat{r}+1} \ge \tilde{\rho} ||E||$  and  $\hat{\lambda}_{\hat{r}+1} \le k \tilde{\rho} ||E||$ . Since  $|\hat{\lambda}_a - \lambda_a| \le ||E||$  for all singular values of X and P, we have  $\lambda_{\hat{r}} - \lambda_{\hat{r}+1} \ge (\tilde{\rho} - 2) ||E||$  and  $\lambda_{\hat{r}+1} \le (k \tilde{\rho} + 1) ||E||$ . Hence,  $\hat{r} \in \mathcal{R}$ .

PROOF OF LEMMA 3.3. From Lemma B.1, we know  $\hat{r}$  exists and  $\hat{r} \in \mathcal{R}$ . Consider an arbitrary  $r \in \mathcal{R}$  and define  $\hat{U}_{1:r} := (\hat{u}_1, \ldots, \hat{u}_r) \in \mathbb{R}^{p \times r}$ . Perform k-means on the columns of  $\hat{U}_{1:r}\hat{U}_{1:r}^T X$  and let the output be

$$\left(\check{z}(r), \left\{\check{\theta}_{j}(r)\right\}_{j=1}^{k}\right) = \underset{z \in [k]^{n}, \left\{\theta_{j}\right\}_{j=1}^{k} \in \mathbb{R}^{p}}{\operatorname{argmin}} \sum_{i \in [n]} \left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}X - \theta_{z_{i}}\right\|^{2}$$

In the following, we are going to establish statistical properties for  $\check{z}(r)$  and eventually obtain a desired upper bound for  $\ell(\check{z}(r), z^*)$ . Since performing k-means on the columns of  $\hat{U}_{1:r}^T X$ is equivalent to k-means on the columns of  $\hat{U}_{1:r} \hat{U}_{1:r}^T X$ , and since  $\hat{r} \in \mathcal{R}$ , we have  $\check{z} = \check{z}(\hat{r})$ and thus the desired upper bound also holds for  $\ell(\check{z}, z^*)$ .

In the rest of the proof we are going to analyze  $\check{z}(r)$  for any  $r \in \mathcal{R}$ . For simplicity, we use the notation  $\check{z}, \{\check{\theta}_j\}_{j \in [n]}$  instead of  $\check{z}(r), \{\check{\theta}_j(r)\}_{j \in [n]}$ . The remaining proof can be decomposed into several parts.

(*Preliminary Results for*  $\check{z}, \{\check{\theta}_j\}_{j\in[n]}$ ). We are going to use Proposition 3.1 to have some preliminary results. Define  $U_{1:r} := (u_1, \ldots, u_r)$  and  $U_{(r+1):k} := (u_{r+1}, \ldots, u_k)$ . Instead of the decomposition (6), we can write

$$X_{i} = U_{1:r}U_{1:r}^{T}\theta_{z_{i}^{*}}^{*} + U_{(r+1):k}U_{(r+1):k}^{T}\theta_{z_{i}^{*}}^{*} + \epsilon_{i} = U_{1:r}U_{1:r}^{T}\theta_{z_{i}^{*}}^{*} + \check{\epsilon}_{i},$$

where  $\check{\epsilon}_i := U_{(r+1):k} U_{(r+1):k}^T \theta_{z_i^*}^* + \epsilon_i$ . In this way, we have a new mixture model with the centers being  $\{U_{1:r}U_{1:r}^T \theta_a^*\}_{a \in [k]}$  and the additive noises being  $\{\check{\epsilon}_i\}$ . Define  $\check{E} := (\check{\epsilon}_1, \ldots, \check{\epsilon}_n)$ . Then

(57)  
$$\begin{split} \|\check{E}\| &\leq \|E\| + \left\| \left( U_{(r+1):k} U_{(r+1):k}^T \theta_{z_1^*}^*, \dots, U_{(r+1):k} U_{(r+1):k}^T \theta_{z_n^*}^* \right) \right\| \\ &= \|E\| + \left\| U_{(r+1):k} U_{(r+1):k}^T P \right\| = \|E\| + \lambda_{r+1} \\ &\leq (k\tilde{\rho} + 2) \|E\|. \end{split}$$

The separation among the new centers is no longer  $\Delta$ . Define

$$\check{\Delta} := \min_{a,b \in [k]: a \neq b} \left\| U_{1:r} U_{1:r}^T \theta_a^* - U_{1:r} U_{1:r}^T \theta_b^* \right\|.$$

For any  $a, b \in [k]$ ,  $U_{1:r}U_{1:r}^T\theta_a^* - U_{1:r}U_{1:r}^T\theta_b^* = (\theta_a^* - \theta_b^*) - U_{(r+1):k}U_{(r+1):k}^T\theta_a^* + U_{(r+1):k}U_{(r+1):k}^T\theta_b^*$ . Also,

$$\max_{a \in [k]} \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_a^* \right\| = \max_{a \in [k]} \sqrt{\frac{\sum_{i \in [n]: z_i^* = a} \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_a^* \right\|^2}{|\{i \in [n]: z_i^* = a\}|}} \le \frac{\left\| U_{(r+1):k} U_{(r+1):k}^T P \right\|_F}{\sqrt{\beta n/k}}$$
(58)
$$\le \frac{2\sqrt{k}\lambda_{r+1}}{\sqrt{\beta n/k}} \le \frac{\sqrt{k}(k\tilde{\rho}+1) \left\| E \right\|}{\sqrt{\beta n/k}}.$$

Hence, we have

(59) 
$$\check{\Delta} \ge \min_{a,b \in [k]: a \neq b} \|\theta_a^* - \theta_b^*\| - 2\max_{a \in [k]} \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_a^* \right\| \ge \Delta - \frac{2\sqrt{k(k\tilde{\rho}+1)} \|E\|}{\sqrt{\beta n/k}}$$

Then from Proposition 3.1, as long as (which will be verified later)

(60) 
$$\check{\psi}_0 := \frac{\Delta}{\beta^{-0.5} k n^{-0.5} \|\check{E}\|} \ge 16,$$

we have

$$\ell(\check{z}, z^*) = \frac{1}{n} |i \in [n] : \check{z}_i \neq \phi(z_i^*)| \le \frac{C_0 k \|\check{E}\|^2}{n\check{\Delta}^2},$$

and

$$\max_{a \in [k]} \left\| \check{\theta}_{\phi(z)} - U_{1:r} U_{1:r}^T \theta_a^* \right\| \le C_0 \beta^{-0.5} k n^{-0.5} \left\| \check{E} \right\|.$$

where  $C_0 = 128$ .

(*Entrywise Decomposition for*  $\check{z}$ ). Next, we are going to have an entrywise decomposition for  $\mathbb{I}\left\{\hat{z}_i \neq \phi(z_i^*)\right\}$  that is analogous to that of Lemma 3.2. When (60) is satisfied, from Lemma 3.1, we have

$$\mathbb{I}\left\{\check{z}_{i}\neq\phi(z_{i}^{*})\right\}\leq\mathbb{I}\left\{\left(1-C_{0}\check{\psi}_{0}^{-1}\right)\check{\Delta}\leq2\left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}\check{\epsilon}_{i}\right\|\right\}.$$

By the definition of  $\check{\epsilon}_i$  and (58), we have

$$\begin{aligned} \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \check{\epsilon}_i \right\| &\leq \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| + \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T U_{(r+1):k} U_{(r+1):k}^T \theta_{z_i^*}^* \right\| \\ &\leq \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| + \left\| U_{(r+1):k} U_{(r+1):k}^T \theta_{z_i^*}^* \right\| \\ &\leq \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| + \frac{\sqrt{k} (k\tilde{\rho} + 1) \left\| E \right\|}{\sqrt{\beta n/k}}. \end{aligned}$$

Then, we have

$$\mathbb{I}\left\{\check{z}_{i}\neq\phi(z_{i}^{*})\right\}\leq\mathbb{I}\left\{\left(1-C_{0}\check{\psi}_{0}^{-1}\right)\check{\Delta}\leq2\left(\left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}\epsilon_{i}\right\|+\frac{\sqrt{k}(k\tilde{\rho}+1)\left\|E\right\|}{\sqrt{\beta n/k}}\right)\right\}$$
$$=\mathbb{I}\left\{\left(1-C_{0}\check{\psi}_{0}^{-1}-\frac{2\sqrt{k}(k\tilde{\rho}+1)\left\|E\right\|}{\sqrt{\beta n/k}\check{\Delta}}\right)\check{\Delta}\leq2\left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}\epsilon_{i}\right\|\right\}.$$

From (56), under the assumption that  $\tilde{\rho} > 4$  and  $\beta n/k^4 > 400$ , we have  $\tilde{\rho}_0$  defined as in (11) to satisfy

$$\tilde{\rho}_0 \ge \frac{(\tilde{\rho} - 1) \|E\|}{\max\left\{\|E\|, \sqrt{\frac{k^2}{\beta n}} (k\tilde{\rho} + 1) \|E\|\right\}} \ge 2.$$

Then Theorem 2.3 can be applied, with which we have

$$\left\| \hat{U}_{1:r} \hat{U}_{1:r}^T - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right\|_{\mathcal{F}} \le \frac{256\sqrt{rk}}{\sqrt{n\beta}} + \frac{256\left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\|}{\lambda_r}.$$

Then following the proof of Lemma 3.2, we have

$$\begin{split} &\mathbb{I}\left\{\check{z}_{i} \neq \phi(z_{i}^{*})\right\} \\ &\leq \mathbb{I}\left\{\left(1 - C_{0}\check{\psi}_{0}^{-1} - \frac{2\sqrt{k}(k\tilde{\rho}+1) \|E\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\| + \left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T} - \hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\right\|_{\mathrm{F}} \|E\|\right)\right\} \\ &\leq \mathbb{I}\left\{\left(1 - C_{0}\check{\psi}_{0}^{-1} - \frac{2\sqrt{k}(k\tilde{\rho}+1) \|E\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(\frac{256\sqrt{rk} \|E\|}{\sqrt{n\beta}} + \left(1 + \frac{256 \|E\|}{\lambda_{r}}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right)\right\} \\ &\leq \mathbb{I}\left\{\left(1 - C_{0}\check{\psi}_{0}^{-1} - \frac{2\sqrt{k}(k\tilde{\rho}+257) \|E\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(1 + \frac{256 \|E\|}{\lambda_{r}}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\} \\ &\leq \mathbb{I}\left\{\left(1 - C_{0}\check{\psi}_{0}^{-1} - \frac{2\sqrt{k}(k\tilde{\rho}+257) \|E\|}{\sqrt{\beta n/k\check{\Delta}}}\right)\check{\Delta} \leq 2\left(1 + \frac{256}{\tilde{\rho}-2}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\}, \end{split}$$

where in the last inequality we use  $\lambda_r \ge (\tilde{\rho} - 2) ||E|| > 0$  (as long as  $\tilde{\rho} > 2$ ) from (56).

The last step of the proof is to simplify the above display using  $\Delta$  instead of  $\dot{\Delta}$ . Then, under the assumption that  $\tilde{\rho} > 256$ , we have  $(1 + 256/(\tilde{\rho} - 2))^{-1} \leq (1 - 512/\tilde{\rho})$ . Recall the definition of  $\tilde{\psi}_0$  in (25). Under the assumption that  $\tilde{\rho} \leq \tilde{\psi}_0/64$ , we have

(61) 
$$\check{\Delta} \ge \Delta \left( 1 - \frac{4\beta^{-0.5}k^2n^{-0.5}\tilde{\rho} \|E\|}{\Delta} \right) = \Delta \left( 1 - \frac{4\tilde{\rho}}{\tilde{\psi}_0} \right) \ge \frac{\Delta}{2}$$

according to (59). Then together with (57), we can verify (60) holds due to

$$\check{\psi}_0 \ge \frac{\Delta/2}{\beta^{-0.5} k n^{-0.5} (k\tilde{\rho}+2) \|E\|} \ge \frac{\Delta}{4\beta^{-0.5} k^2 n^{-0.5} \tilde{\rho} \|E\|} = \frac{\psi_0}{4\tilde{\rho}} \ge 16$$

Rearranging all the terms with the help of (61), we can simplify  $\mathbb{I}\{\check{z}_i \neq \phi(z_i^*)\}$  into

$$\mathbb{I}\left\{\tilde{z}_{i} \neq \phi(z_{i}^{*})\right\}$$

$$\leq \mathbb{I}\left\{\left(1 - 4C_{0}\tilde{\rho}\tilde{\psi}_{0} - \frac{4\beta^{-0.5}k^{2}n^{-0.5}\tilde{\rho} \|E\|}{\Delta/2}\right)\left(1 - \frac{256}{\tilde{\rho}}\right)\left(1 - \frac{4\tilde{\rho}}{\tilde{\psi}_{0}}\right)\Delta \leq 2\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\}$$

$$\leq \mathbb{I}\left\{\left(1 - 5C_{0}\tilde{\rho}\tilde{\psi}_{0}^{-1} - 256\tilde{\rho}^{-1}\right)\Delta \leq 2\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\}.$$

PROOF OF THEOREM 3.2. Recall the definition of  $\mathcal{F}$  in (46). Then if  $\mathcal{F}$  holds, by appropriate choices of  $C_1, C_2$ , we can verify the assumptions needed in Lemma 3.3 hold, which lead to

$$\mathbb{I}\{\tilde{z}_{i} \neq \phi(z_{i}^{*})\}\mathbb{I}\{\mathcal{F}\} \leq \mathbb{I}\left\{\left(1 - C''(\rho_{2}\psi_{2}^{-1} + \rho_{2}^{-1})\right)\Delta \leq 2\left\|\hat{U}_{-i,1:\hat{r}}\hat{U}_{-i,1:\hat{r}}^{T}\epsilon_{i}\right\|\right\}\mathbb{I}\{\mathcal{F}\},\$$

for some constant C'' > 0. Though  $\hat{r}$  is random, the proof of Lemma 3.3 shows that  $\hat{r} \in \mathcal{R} \subset [k]$  where  $\mathcal{R}$  is defined in (56). Note that for any  $r \in [k]$ , we can follow the proof of Theorem 3.1 to show

$$\mathbb{EI}\left\{\left(1-C''(\rho_{2}\psi_{2}^{-1}+\rho_{2}^{-1})\right)\Delta\leq 2\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\}\leq\exp\left(-(1-C'''(\rho_{2}\psi_{2}^{-1}+\rho_{2}^{-1}))\frac{\Delta^{2}}{8\sigma^{2}}\right),$$

for some constant C''' > 0. Hence, the same upper bound holds for  $\mathbb{EI}\{(1 - C''(\rho_2\psi_2^{-1} + \rho_2^{-1}))\Delta \le 2\|\hat{U}_{-i,1:\hat{r}}\hat{U}_{-i,1:\hat{r}}^T\epsilon_i\|\}$ . The rest of the proof follows that of Theorem 3.1 and is omitted here.

#### APPENDIX C: PROOF OF THEOREM 3.3

Define  $\mathcal{F} = \{ \|E\| \le \sqrt{2}(\sqrt{n} + \sqrt{p})\sigma \}$ . Then by Lemma B.1 of [27], we have  $\mathbb{P}(\mathcal{F}) \ge 1 - e^{-0.08n}$ . Then under the event  $\mathcal{F}$ , the assumption (26) implies (16) holds, and hence (17) and (18) hold. For simplicity, and without loss of generality, we can let  $\phi$  in (17)-(18) to be the identity, and we get

$$\ell(\hat{z}, z^*) = \frac{1}{n} |\{i \in [n] : \hat{z}_i \neq z_i^*\}| \le \frac{C_0 k \left(1 + \sqrt{\frac{p}{n}}\right)^2 \sigma^2}{\Delta^2},$$

and

$$\max_{a \in [k]} \left\| \hat{\theta}_a - \theta_a^* \right\| \le C_0 \beta^{-0.5} k \left( 1 + \sqrt{\frac{p}{n}} \right) \sigma,$$

where  $C_0 > 0$  is some constant.

Denote  $\hat{P} = \hat{U}_{1:k} \hat{U}_{1:k}^T X$  and let  $\hat{P}_{\cdot,i}$  be its *i*th column so that  $\hat{P}_{\cdot,i} = \hat{U}_{1:k} \hat{U}_{1:k}^T X_i$ . We define  $r \in [k]$  as (with  $\lambda_{k+1} := 0$ )

(62) 
$$r = \max\left\{j \in [k] : \lambda_j - \lambda_{j+1} \ge \tau \sqrt{n+p\sigma}\right\},$$

for a sequence  $\tau \to \infty$  to be determined later. We note that if  $\Delta/(k^{\frac{3}{2}}\tau\beta^{\frac{1}{2}}(1+p/n)^{\frac{1}{2}}\sigma) \to \infty$ , the set  $\{j \in [k] : \lambda_j - \lambda_{j+1} \ge \tau\sqrt{n+p}\sigma\}$  is not empty. Otherwise, this would imply  $\lambda_1 \le k\tau\sqrt{n+p}\sigma$  which would contradict with the fact  $\lambda_1 \ge \sqrt{\beta n/k}\Delta/(2\sigma)$  (see Proposition A.1 of [27]). By the definition of r in (62), we immediately have

(63) 
$$\lambda_r - \lambda_{r+1} \ge \tau \sqrt{n+p}\sigma,$$

(64) and 
$$\lambda_{r+1} \le k\tau \sqrt{n+p}\sigma$$
.

We split  $\hat{U}_{1:k}$  into  $(\hat{U}_{1:r}, \hat{U}_{(r+1):k})$  where  $\hat{U}_{1:r} := (\hat{u}_1, \dots, \hat{u}_r)$  and  $\hat{U}_{(r+1):k} := (\hat{u}_{r+1}, \dots, \hat{u}_k)$ . We decompose  $\hat{P}_{\cdot,i} = \hat{P}_{\cdot,i}^{(1)} + \hat{P}_{\cdot,i}^{(2)}$ , where  $\hat{P}_{\cdot,i}^{(1)} := \hat{U}_{1:r}\hat{U}_{1:r}^T\hat{P}_{\cdot,i}$  and  $\hat{P}_{\cdot,i}^{(2)} := \hat{U}_{(r+1):k}\hat{U}_{(r+1):k}^T\hat{P}_{\cdot,i}$ . Similarly, for each  $a \in [k]$ , we decompose  $\hat{\theta}_a = \hat{\theta}_a^{(1)} + \hat{\theta}_a^{(2)}$ , where  $\hat{\theta}_a^{(1)} := \hat{U}_{1:r}\hat{U}_{1:r}^T\hat{\theta}_a$  and  $\hat{\theta}_a^{(2)} := \hat{U}_{(r+1):k}\hat{U}_{(r+1):k}^T\hat{\theta}_a$ . Due to the orthogonality of  $\{\hat{u}_l\}_{l\in[k]}$ , we obtain that for any  $i \in [n]$  and any  $a \in [k]$  such that  $a \neq z_i^*$ ,

$$\mathbb{I}\left\{\hat{z}_{i}=a\right\} \leq \mathbb{I}\left\{\left\|\hat{P}_{\cdot,i}^{(1)}+\hat{P}_{\cdot,i}^{(2)}-\hat{\theta}_{a}^{(1)}-\hat{\theta}_{a}^{(2)}\right\|^{2} \leq \left\|\hat{P}_{\cdot,i}^{(1)}+\hat{P}_{\cdot,i}^{(2)}-\hat{\theta}_{z_{i}^{*}}^{(1)}-\hat{\theta}_{z_{i}^{*}}^{(2)}\right\|^{2}\right\} \\
= \mathbb{I}\left\{2\left\langle\hat{P}_{\cdot,i}^{(1)}-\hat{\theta}_{z_{i}^{*}}^{(1)},\hat{\theta}_{z_{i}^{*}}^{(1)}-\hat{\theta}_{a}^{(1)}\right\rangle+\left\|\hat{\theta}_{z_{i}^{*}}^{(1)}-\hat{\theta}_{a}^{(1)}\right\|^{2} \leq 2\left\langle\hat{P}_{\cdot,i}^{(2)},\hat{\theta}_{a}^{(2)}-\hat{\theta}_{z_{i}^{*}}^{(2)}\right\rangle-\left\|\hat{\theta}_{a}^{(2)}\right\|^{2}+\left\|\hat{\theta}_{z_{i}^{*}}^{(2)}\right\|^{2}\right\}$$

We denote  $\tau'' = o(1)$  to be another sequence which we will specify later. Then the above display can be decomposed and upper bounded by

$$\mathbb{I}\left\{\hat{z}_{i}=a\right\} \leq \mathbb{I}\left\{\left\|\hat{\theta}_{z_{i}^{*}}^{(1)}-\hat{\theta}_{a}^{(1)}\right\|-\frac{\tau''\Delta^{2}+\left\|\hat{\theta}_{z_{i}^{*}}^{(2)}\right\|^{2}}{\left\|\hat{\theta}_{z_{i}^{*}}^{(1)}-\hat{\theta}_{a}^{(1)}\right\|}\leq 2\left\|\hat{P}_{\cdot,i}^{(1)}-\hat{\theta}_{z_{i}^{*}}^{(1)}\right\|\right\}$$
$$+\mathbb{I}\left\{\tau''\Delta^{2}\leq 2\left\langle\hat{P}_{\cdot,i}^{(2)},\hat{\theta}_{a}^{(2)}-\hat{\theta}_{z_{i}^{*}}^{(2)}\right\rangle\right\}=:A_{i,a}+B_{i,a}.$$

Then

$$\mathbb{E}\ell(\hat{z}, z^*) \leq \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}\mathbb{I}\{\hat{z}_i = a\}$$
(65) 
$$\leq \mathbb{P}\left(\mathcal{F}^{\complement}\right) + \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}A_{i,a}\mathbb{I}\{\mathcal{F}\} + \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}B_{i,a}\mathbb{I}\{\mathcal{F}\}.$$

We are going to establish upper bounds first for  $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}B_{i,a}\mathbb{I}\{\mathcal{F}\}$  and then for  $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}A_{i,a}\mathbb{I}\{\mathcal{F}\}.$ 

(Analysis on  $n^{-1}\sum_{i\in[n]}\sum_{a\neq z_i^*}\mathbb{E}B_{i,a}\mathbb{I}\{\mathcal{F}\}$ ). For  $\sum_{i\in[n]}\sum_{a\neq z_i^*}\mathbb{E}B_{i,a}\mathbb{I}\{\mathcal{F}\}$ , we can di-

rectly use upper bounds established in Section 4.4.3 of [27]<sup>1</sup>. It proves that for any  $i \in [n]$ ,

$$\sum_{a \in [k]: a \neq z_i^*} B_{i,a} \mathbb{I}\left\{\mathcal{F} \cap \mathcal{T}\right\} \le 2 \exp\left(-\frac{1}{2} \left(c_4 \frac{\tau'' \Delta}{k^{\frac{7}{2}} \tau^2 \beta^{-\frac{1}{2}} (1+\frac{p}{n}) \sigma} \sqrt{\frac{n-k}{3n}}\right)^2 \frac{\Delta^2}{\sigma^2}\right),$$

where  $c_4 > 0$  is some constant, and T is some high-probability event in the sense that

$$\mathbb{P}(\mathcal{T}) \ge 1 - nk \exp\left(-\frac{(n-k)}{9}\right).$$

Hence,

$$\begin{aligned} \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}B_{i,a} \mathbb{I}\left\{\mathcal{F}\right\} &\leq \frac{1}{n} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}B_{i,a} \mathbb{I}\left\{\mathcal{F} \cap \mathcal{T}\right\} + \mathbb{P}\left(\mathcal{T}^{\complement}\right) \\ &\leq 2 \exp\left(-\frac{1}{2} \left(c_4 \frac{\tau'' \Delta}{k^{\frac{7}{2}} \tau^2 \beta^{-\frac{1}{2}} (1 + \frac{p}{n}) \sigma} \sqrt{\frac{n-k}{3n}}\right)^2 \frac{\Delta^2}{\sigma^2}\right) + nk \exp\left(-\frac{(n-k)}{9}\right). \end{aligned}$$

(Analysis on  $n^{-1}\sum_{i\in[n]}\sum_{a\neq z_i^*} \mathbb{E}A_{i,a}\mathbb{I}\{\mathcal{F}\}$ ). We first follow some algebra as in Section 4.4.2 of [27] to simplify  $A_{i,a}\mathbb{I}\{\mathcal{F}\}$ . For any  $i\in[n]$  and  $a\neq z_i^*$ , it proves

(66) 
$$A_{i,a}\mathbb{I}\left\{\mathcal{F}\right\} \leq \mathbb{I}\left\{\left(1 - c_{1}\tau'' - \frac{c_{1}k^{2}\tau\beta^{-\frac{1}{2}}\sqrt{1 + \frac{p}{n}}\sigma}{\Delta}\right)\Delta \leq 2\left\|\hat{P}_{\cdot,i}^{(1)} - \hat{\theta}_{z_{i}^{*}}^{(1)}\right\|\right\}\mathbb{I}\left\{\mathcal{F}\right\},$$

for some constant  $c_1 > 0$ . Still working on the event  $\mathcal{F}$ , it also proves

(67) 
$$\left\| \hat{P}_{\cdot,i}^{(1)} - \hat{\theta}_{z_i^*}^{(1)} \right\| \le \left\| \hat{P}_{\cdot,i}^{(1)} - \hat{U}_{1:r} \hat{U}_{1:r}^T \theta_{z_i^*}^* \right\| + 8\sqrt{2}\sqrt{\beta^{-1}k^2 \left(1 + \frac{p}{n}\right)\sigma}.$$

Our following analysis on  $A_{i,a}\mathbb{I}\{\mathcal{F}\}$  is different from the rest proof in Section 4.4.2 of [27]. Note that  $\hat{P}_{:,i}^{(1)} - \hat{U}_{1:r}\hat{U}_{1:r}^T\theta_{z_i^*} = \hat{U}_{1:r}\hat{U}_{1:r}^TX_i - \hat{U}_{1:r}\hat{U}_{1:r}^T\theta_{z_i^*} = \hat{U}_{1:r}\hat{U}_{1:r}^T\epsilon_i$ . Then (66) and (67) give

(68) 
$$A_{i,a}\mathbb{I}\left\{\mathcal{F}\right\} \leq \mathbb{I}\left\{\left(1 - c_{2}\tau'' - \frac{c_{2}k^{2}\tau\beta^{-\frac{1}{2}}\left(1 + \sqrt{\frac{p}{n}}\right)\sigma}{\Delta}\right)\Delta \leq 2\left\|\hat{U}_{1:r}\hat{U}_{1:r}^{T}\epsilon_{i}\right\|\right\}\mathbb{I}\left\{\mathcal{F}\right\},$$

where we use  $\tau \to \infty$  and the fact that  $1 + \sqrt{p/n}, \sqrt{1 + p/n}$  are of the same order.

Recall the definition of  $X_{-i}$  in (8) and  $\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T$  is the leave-one-out counterpart of  $\hat{U}_{1:r}\hat{U}_{1:r}^T$ . For (68), we can decompose  $\|\hat{U}_{1:r}\hat{U}_{1:r}^T\epsilon_i\|$  into

$$\left\| \hat{U}_{1:r} \hat{U}_{1:r}^T \epsilon_i \right\| \le \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\| + \left\| \hat{U}_{1:r} \hat{U}_{1:r}^T - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right\|_{\mathbf{F}} \|\epsilon_i\|.$$

To upper bound  $\|\hat{U}_{1:r}\hat{U}_{1:r}^T - \hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\|_F$ , we are going to use Theorem 2.3. Since (63)-(64) hold, under the assumption  $\beta n/k^4 \ge 100$ , we have

$$\frac{\lambda_r - \lambda_{r+1}}{\max\left\{\|E\|, \sqrt{\frac{k^2}{n\beta}}\lambda_{r+1}\right\}} \ge \frac{\tau}{2}.$$

<sup>&</sup>lt;sup>1</sup>The model in [27] assumes  $\{\epsilon_j\} \stackrel{iid}{\sim} \mathcal{N}(0, I)$  while in this paper we assume  $\{\epsilon_j\} \stackrel{iid}{\sim} \mathcal{N}(0, \sigma^2 I)$ . To directly use results from [27], we can re-scale our data to have  $X'_j = X_j / \sigma$  for all  $j \in [n]$ . Then  $\{X'_j\}$  has  $\mathcal{N}(0, I)$  noise and the separation between their centers becomes  $\Delta / \sigma$ . Then all the results from [27] can be used here with  $\Delta$  replaced by  $\Delta / \sigma$ .

Applying Theorem 2.3, we have

$$\left| \hat{U}_{1:r} \hat{U}_{1:r}^T - \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \right\|_{\mathcal{F}} \le \frac{256\sqrt{rk}}{\sqrt{n\beta}} + \frac{256 \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^T \epsilon_i \right\|}{\lambda_r}.$$

Hence,

$$\begin{split} \left\| \hat{U}_{1:r} \hat{U}_{1:r}^{T} \epsilon_{i} \right\| &\leq \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^{T} \epsilon_{i} \right\| + \left( \frac{256\sqrt{rk}}{\sqrt{n\beta}} + \frac{256 \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^{T} \epsilon_{i} \right\|}{\lambda_{r}} \right) \| E \| \\ &= \frac{256k \left\| E \right\|}{\sqrt{n\beta}} + \left( 1 + \frac{256 \left\| E \right\|}{\lambda_{r}} \right) \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^{T} \epsilon_{i} \right\| \\ &\leq \frac{256\sqrt{2}k(\sqrt{n} + \sqrt{p})\sigma}{\sqrt{n\beta}} + \left( 1 + \frac{256\sqrt{2}(\sqrt{n} + \sqrt{p})\sigma}{\tau\sqrt{n + p\sigma}} \right) \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^{T} \epsilon_{i} \right\| \\ &\leq 512k\beta^{-0.5} \left( 1 + \sqrt{\frac{p}{n}} \right) \sigma + \left( 1 + 512\tau^{-1} \right) \left\| \hat{U}_{-i,1:r} \hat{U}_{-i,1:r}^{T} \epsilon_{i} \right\|, \end{split}$$

where in the second to the last inequality, we use (63) for  $\lambda_r$  and the event  $\mathcal{F}$  for ||E||. Then (68) leads to

$$\begin{aligned} A_{i,a}\mathbb{I}\left\{\mathcal{F}\right\} &\leq \mathbb{I}\left\{\left(1-c_{3}\tau''-\frac{c_{3}k^{2}\tau\beta^{-\frac{1}{2}}\left(1+\sqrt{\frac{p}{n}}\right)\sigma}{\Delta}\right)\Delta \leq 2\left(1+512\tau^{-1}\right)\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\}\mathbb{I}\left\{\mathcal{F}\right\} \\ &\leq \mathbb{I}\left\{\left(1-c_{4}\left(\frac{k^{2}\tau\beta^{-\frac{1}{2}}\left(1+\sqrt{\frac{p}{n}}\right)\sigma}{\Delta}+\tau^{-1}\right)\right)\Delta \leq 2\left\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^{T}\epsilon_{i}\right\|\right\},\end{aligned}$$

where  $c_3, c_4 > 0$  are some constants. As long as  $1 - c_4(k^2\tau\beta^{-0.5}(1+\sqrt{p/n})\sigma/\Delta+\tau^{-1}) > 1/2$ , we can use Lemma E.2 to calculate the tail probability of  $\|\hat{U}_{-i,1:r}\hat{U}_{-i,1:r}^T\epsilon_i\|$ . Following the proof of Theorem 3.1, we have

$$\mathbb{E}A_{i,a}\mathbb{I}\left\{\mathcal{F}\right\} \leq \exp\left(-\left(1-c_5\left(\frac{k^2\tau\beta^{-\frac{1}{2}}\left(1+\sqrt{\frac{p}{n}}\right)\sigma}{\Delta}+\tau^{-1}\right)\right)\frac{\Delta^2}{8\sigma^2}\right),$$

for some constant  $c_5 > 0$ . Then we have,

$$n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}A_{i,a} \mathbb{I}\left\{\mathcal{F}\right\} \le k \exp\left(-\left(1 - c_5\left(\frac{k^2 \tau \beta^{-\frac{1}{2}} \left(1 + \sqrt{\frac{p}{n}}\right)\sigma}{\Delta} + \tau^{-1}\right)\right) \frac{\Delta^2}{8\sigma^2}\right).$$

(Obtaining the Final Result.) From (65) and the above upper bounds on  $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}B_{i,a}\mathbb{I}\{\mathcal{F}\}$ and  $n^{-1} \sum_{i \in [n]} \sum_{a \in [k]: a \neq z_i^*} \mathbb{E}A_{i,a}\mathbb{I}\{\mathcal{F}\}$ , we have

$$\mathbb{E}\ell(\hat{z}, z^*) \le e^{-0.08n} + 2\exp\left(-\frac{1}{2}\left(c_4 \frac{\tau''\Delta}{k^{\frac{7}{2}}\tau^2\beta^{-\frac{1}{2}}(1+\frac{p}{n})\sigma}\sqrt{\frac{n-k}{3n}}\right)^2 \frac{\Delta^2}{\sigma^2}\right) + nk\exp\left(-\frac{(n-k)}{9}\right) + k\exp\left(-\left(1-c_5 \left(\frac{k^2\tau\beta^{-\frac{1}{2}}\left(1+\sqrt{\frac{p}{n}}\right)\sigma}{\Delta} + \tau^{-1}\right)\right)\frac{\Delta^2}{8\sigma^2}\right).$$

Since we assume  $\beta n/k^4 \ge 100$ , we have (n-k)/n > 0.99. Hence, under the assumption that  $\Delta/(k^{3.5}\beta^{-0.5}(1+\frac{p}{n})\sigma) \to \infty$ , we can take  $\tau, \tau''$  to be

$$\tau = \tau^{\prime\prime - 1} := \left(\frac{\Delta}{k^{3.5}\beta^{-0.5} \left(1 + \frac{p}{n}\right)\sigma}\right)^{0.25}$$

such that  $\tau \to \infty$  and  $\tau'' = o(1)$ . Then for some constant  $c_6 > 0$ , we have

$$\begin{split} \mathbb{E}\ell(\hat{z}, z^*) &\leq e^{-0.08n} + 2\exp\left(-\frac{c_4^2}{12} \left(\frac{\Delta}{k^{3.5}\beta^{-0.5} \left(1+\frac{p}{n}\right)\sigma}\right)^{0.5} \frac{\Delta^2}{\sigma^2}\right) + nke^{-0.1n} \\ &+ k\exp\left(-\left(1 - 2c_5 \left(\frac{\Delta}{k^{3.5}\beta^{-0.5} \left(1+\frac{p}{n}\right)\sigma}\right)^{-0.25}\right) \frac{\Delta^2}{8\sigma^2}\right) \\ &\leq \exp\left(-\left(1 - c_6 \left(\frac{\Delta}{k^{3.5}\beta^{-0.5} \left(1+\frac{p}{n}\right)\sigma}\right)^{-0.25}\right) \frac{\Delta^2}{8\sigma^2}\right) + 2e^{-0.08n}. \end{split}$$

## APPENDIX D: PROOFS OF RESULTS IN SECTION 3.6

**D.1. Proof of Theorem 3.4.** The proof of Theorem 3.4 relies on the following entrywise decomposition that is analogous to Lemma 3.2 but in an opposite direction. Note the the singular vectors  $\hat{u}_1$ , and  $\{\hat{u}_{1,-i}\}_{i\in[n]}$  are all identifiable up to sign. Without loss of generality, we assume  $\langle \hat{u}_1, u_1 \rangle \ge 0$  and  $\langle \hat{u}_{1,-i}, u_1 \rangle \ge 0$  for all  $i \in [n]$ .

LEMMA D.1. Consider the model (28). Let  $\phi \in \Phi$  be the permutation such that  $\ell(\check{z}, z^*) = \frac{1}{n} |\{i \in [n] : \check{z}_i \neq \phi(z_i^*)\}|$ . Then there exists a constants  $C, C_1 > 0$  such that if

(69) 
$$\frac{\Delta}{\beta^{-0.5}n^{-0.5} \|E\|} \ge C,$$

then for any  $i \in [n]$ ,

(70) 
$$\mathbb{I}\left\{\check{z}_{i}\neq\phi(z_{i}^{*})\right\}\geq\mathbb{I}\left\{\left(1+\frac{C_{1}\beta^{-0.5}n^{-0.5}\|E\|}{\Delta}\right)\Delta\leq-2(\hat{u}_{1,-i}^{T}\epsilon_{i})sign(u_{1}^{T}\theta_{\phi(z_{i}^{*})})\right\}.$$

PROOF. The proof mainly follows the proofs of Lemma 3.1 and Lemma 3.2 with some modifications such as adding a negative term instead of a positive term in order to obtain a lower bound.

We first write  $\check{z}$  equivalently as

$$\left(\check{z}, \left\{\check{\theta}_{j}\right\}_{j=1}^{2}\right) = \operatorname*{argmin}_{z \in [2]^{n}, \left\{\theta_{j}\right\}_{j=1}^{2} \in \mathbb{R}^{p}} \sum_{i \in [n]} \left\|\hat{u}_{1}\hat{u}_{1}^{T}X_{i} - \theta_{z_{i}}\right\|^{2},$$

where  $\dot{\theta}_a = \hat{u}_1 \check{c}_a$  for each  $a \in [2]$ . Note that k = 2. From Proposition 3.1, we have

$$\frac{1}{n} |\{i \in [n] : \check{z}_i \neq \phi(z_i^*)\}| \le \frac{C_0 k ||E||^2}{n\Delta^2},$$

and

(71) 
$$\max_{a \in [2]} \left\| \check{\theta}_{\phi(a)} - \theta_a^* \right\| \le C_0 \beta^{-0.5} k n^{-0.5} \left\| E \right\|,$$

for some permutation  $\phi : [2] \to [2]$  and some constant  $C_0 > 0$ . Without loss of generality, assume  $\phi = \text{Id}$ .

Recall that  $\theta_1^* = -\theta_2^* = \delta \mathbb{1}_p$ ,  $u_1 = 1/\sqrt{p}\mathbb{1}_p$ ,  $\lambda_1 = \delta\sqrt{np} = \frac{\Delta\sqrt{n}}{2}$ , and  $|u_1^T(\theta_{z_i^*}^* - (-\theta_{z_i^*}^*))| = 2\delta\sqrt{p} = \Delta$ . By Davis-Kahan Theorem, we have

$$\min_{s \in \pm 1} \|\hat{u}_1 - su_1\| \le \frac{\|E\|}{\lambda_1} = \frac{2\|E\|}{\sqrt{n\Delta}} \le 1/16,$$

where the last inequality is due to the assumption (16). Since we assume  $\langle \hat{u}_1, u_1 \rangle \ge 0$ , we have  $\|\hat{u}_1 - su_1\| = \min_{s \in \pm 1} \|\hat{u}_1 - su_1\|$ .

Consider any  $i \in [n]$  and any  $a \in [2]$  such that  $a \neq z_i^*$ . Note that for any scalars x, y, w, if  $|x - y| \leq |x - w|$ , we have equivalently  $\operatorname{sign}(w - y)(y + w)/2 \geq \operatorname{sign}(w - y)x$ . Since (y + w)/2 = (y - w)/2 + w, a sufficient condition is  $|w - y|/2 + |w| \leq (-\operatorname{sign}(w - y))x$ . Hence, we have

$$\begin{split} & \mathbb{I}\left\{ \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \dot{\theta}_{a} \right\| \leq \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \dot{\theta}_{z_{i}^{*}} \right\| \right\} \\ &= \mathbb{I}\left\{ \left| \hat{u}_{1}^{T} X_{i} - \hat{u}_{1}^{T} \check{\theta}_{a} \right| \leq \left| \hat{u}_{1}^{T} X_{i} - \hat{u}_{1}^{T} \check{\theta}_{z_{i}^{*}} \right| \right\} \\ &= \mathbb{I}\left\{ \left| \hat{u}_{1}^{T} \epsilon_{i} - \hat{u}_{1}^{T} \left( \check{\theta}_{a} - \theta_{z_{i}^{*}}^{*} \right) \right| \leq \left| \hat{u}_{1}^{T} \epsilon_{i} - \hat{u}_{1}^{T} \left( \check{\theta}_{z_{i}^{*}} - \theta_{z_{i}^{*}}^{*} \right) \right| \right\} \\ &\geq \mathbb{I}\left\{ \frac{1}{2} \left| \hat{u}_{1}^{T} (\check{\theta}_{z_{i}^{*}} - \check{\theta}_{a}) \right| + \left| \hat{u}_{1}^{T} \left( \check{\theta}_{z_{i}^{*}} - \theta_{z_{i}^{*}}^{*} \right) \right| \leq -(\hat{u}_{1}^{T} \epsilon_{i}) \mathrm{sign}(\hat{u}_{1}^{T} (\check{\theta}_{z_{i}^{*}} - \check{\theta}_{a})) \right\} \\ &\geq \mathbb{I}\left\{ \left\| \check{\theta}_{z_{i}^{*}} - \check{\theta}_{a} \right\| + 2 \left\| \check{\theta}_{z_{i}^{*}} - \theta_{z_{i}^{*}}^{*} \right\| \leq -2(\hat{u}_{1}^{T} \epsilon_{i}) \mathrm{sign}(\hat{u}_{1}^{T} (\check{\theta}_{z_{i}^{*}} - \check{\theta}_{a})) \right\}. \end{split}$$

We are going to show  $\operatorname{sign}(\hat{u}_1^T(\check{\theta}_{z_i^*} - \check{\theta}_a)) = \operatorname{sign}(u_1^T(\theta_{z_i^*}^* - \theta_a^*))$ . By (71), we have

$$\begin{split} \left\langle \check{\theta}_{z_i^*} - \check{\theta}_a, \theta_{z_i^*}^* - \theta_a^* \right\rangle &= \left\| \theta_{z_i^*}^* - \theta_a^* \right\|^2 + \left\langle \check{\theta}_{z_i^*} - \theta_{z_i^*}^*, \theta_{z_i^*}^* - \theta_a^* \right\rangle + \left\langle \check{\theta}_a - \theta_a^*, \theta_{z_i^*}^* - \theta_a^* \right\rangle \\ &\geq \Delta^2 \left( 1 - \frac{2C_0 k \beta^{-0.5} n^{-0.5} \|E\|}{\Delta} \right) \\ &> 0, \end{split}$$

where the last inequality holds as long as  $\Delta > 2C_0\beta^{-0.5}kn^{-0.5}||E||$ . Due to the fact  $\theta_{z_i^*}^* - \theta_a^* \in \operatorname{span}(u_1)$ ,  $\check{\theta}_{z_i^*} - \check{\theta}_a^* \in \operatorname{span}(\hat{u}_1)$ , and  $\langle \hat{u}_1, u_1 \rangle \ge 0$ , if  $u_1, \theta_{z_i^*}^* - \theta_a^*$  are in the same direction, then  $\hat{u}_1, \check{\theta}_{z_i^*} - \check{\theta}_a^*$  must also be in the same direction, and vice versa. Hence, we have  $\operatorname{sign}(\hat{u}_1^T(\check{\theta}_{z_i^*} - \check{\theta}_a)) = \operatorname{sign}(u_1^T(\theta_{z_i^*}^* - \theta_a^*))$ . Thus,

$$\mathbb{I}\left\{ \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \check{\theta}_{a} \right\| \leq \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \check{\theta}_{z_{i}^{*}} \right\| \right\}$$

$$\geq \mathbb{I}\left\{ \left\| \check{\theta}_{z_{i}^{*}} - \check{\theta}_{a} \right\| + 2 \left\| \check{\theta}_{z_{i}^{*}} - \theta_{z_{i}^{*}}^{*} \right\| \leq -2(\hat{u}_{1}^{T} \epsilon_{i}) \operatorname{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*} - \theta_{a}^{*})) \right\}.$$

Following the same analysis as in the proof of Lemma 3.1, we can get the following result that is analogous to (45):

$$\mathbb{I}\left\{ \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \check{\theta}_{a} \right\| \leq \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \check{\theta}_{z_{i}^{*}} \right\| \right\}$$

$$\geq \mathbb{I}\left\{ \left( 1 + \frac{4C_{0}\beta^{-0.5}kn^{-0.5} \|E\|}{\Delta} \right) \Delta \leq -2(\hat{u}_{1}^{T}\epsilon_{i}) \operatorname{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*} - \theta_{a}^{*})) \right\}$$

Next, we are going to decompose  $\hat{u}_1^T \epsilon_i$  following the proof of Lemma 3.2. Denote  $\hat{u}_{1,-i}$  be the leave-one-out counterpart of  $\hat{u}_1$ , i.e.,  $\hat{u}_{1,-i}$  is the leading left singular vector of  $X_{-i}$ .

Since we assume  $\langle \hat{u}_{1,-i}, u_1 \rangle \ge 0$ , we have  $\|\hat{u}_{1,-i} - u_1\| \le 2 \|E\| / (\sqrt{n-1}\Delta)$ . As a result, we have  $\|\hat{u}_{1,-i} - \hat{u}_1\| \le 4 \|E\| / (\sqrt{n-1}\Delta)$  which leads to

(72) 
$$\langle \hat{u}_{1,-i}, \hat{u}_1 \rangle \ge 1 - 4 \|E\| / (\sqrt{n-1}\Delta) > 0.$$

We have the following decomposition:

$$\begin{split} &(\hat{u}_{1}^{T}\epsilon_{i})\mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\ &= \left\langle \hat{u}_{1}, \hat{u}_{1}\hat{u}_{1}^{T}\epsilon_{i} \right\rangle \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\ &= \left\langle \hat{u}_{1}, (\hat{u}_{1,-i}\hat{u}_{1,-i}^{T})\epsilon_{i} \right\rangle \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \left\langle \hat{u}_{1}, (\hat{u}_{1}\hat{u}_{1}^{T}-\hat{u}_{1,-i}\hat{u}_{1,-i}^{T})\epsilon_{i} \right\rangle \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\ &= \left\langle \hat{u}_{1}, \hat{u}_{1,-i} \right\rangle (\hat{u}_{1,-i}^{T}\epsilon_{i})\mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \left\langle \hat{u}_{1}, (\hat{u}_{1}\hat{u}_{1}^{T}-\hat{u}_{1,-i}\hat{u}_{1,-i}^{T})\epsilon_{i} \right\rangle \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\ &\leq \left\langle \hat{u}_{1}, \hat{u}_{1,-i} \right\rangle (\hat{u}_{1,-i}^{T}\epsilon_{i})\mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \left\| \hat{u}_{1}\hat{u}_{1}^{T}-\hat{u}_{1,-i}\hat{u}_{1,-i}^{T} \right\| \left\| \epsilon_{i} \right\|. \end{split}$$

Note that  $\lambda_1 / ||E|| = \Delta \sqrt{n} / (2 ||E||)$  is greater than 16 under the assumption (69) holds for a large constant *C*. From Theorem 2.2 we have

$$\left\| \hat{u}_{1} \hat{u}_{1}^{T} - \hat{u}_{1,-i} \hat{u}_{1,-i}^{T} \right\| \leq \frac{128}{\lambda_{1} / \|E\|} \left( \frac{k}{\sqrt{\beta n}} + \frac{\left\| \hat{u}_{1,-i} \hat{u}_{1,-i}^{T} \epsilon_{i} \right\|}{\lambda_{1}} \right).$$

Then,

$$\begin{split} &(\hat{u}_{1}^{T}\epsilon_{i})\mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) \\ &\leq \langle \hat{u}_{1}, \hat{u}_{1,-i}\rangle \left( \hat{u}_{1,-i}^{T}\epsilon_{i} \right) \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \left( \frac{128k}{\sqrt{n\beta}(\lambda_{1}/\|E\|)} + \frac{128\left\| \hat{u}_{1,-i}\hat{u}_{1,-i}^{T}\epsilon_{i} \right\|}{\lambda_{1}^{2}/\|E\|} \right) \|E\| \\ &= \langle \hat{u}_{1}, \hat{u}_{1,-i}\rangle \left( \hat{u}_{1,-i}^{T}\epsilon_{i} \right) \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*}-\theta_{a}^{*})) + \frac{256n^{-0.5}k\beta^{-0.5}\|E\|^{2}}{\Delta} + \frac{512\left| \hat{u}_{1,-i}^{T}\epsilon_{i} \right| n^{-1}\|E\|^{2}}{\Delta^{2}} \right. \\ &\text{So far we have obtained} \\ &\mathbb{I}\left\{ \left\| \hat{u}_{1}\hat{u}_{1}^{T}X_{i} - \check{\theta}_{a} \right\| \leq \left\| \hat{u}_{1}\hat{u}_{1}^{T}X_{i} - \check{\theta}_{z_{i}^{*}} \right\| \right\} \\ &\geq \mathbb{I}\left\{ \left( 1 + \frac{4C_{0}\beta^{-0.5}kn^{-0.5}\|E\|}{\Delta} \right) \Delta \leq -2\langle \hat{u}_{1}, \hat{u}_{1,-i}\rangle \left( \hat{u}_{1,-i}^{T}\epsilon_{i} \right) \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*} - \theta_{a}^{*})) \right. \\ &\left. - \frac{256n^{-0.5}k\beta^{-0.5}\|E\|^{2}}{\Delta} - \frac{512\left| \hat{u}_{1,-i}^{T}\epsilon_{i} \right| n^{-1}\|E\|^{2}}{\Delta^{2}} \right\} \\ &= \mathbb{I}\left\{ \left( 1 + \frac{4C_{0}\beta^{-0.5}kn^{-0.5}\|E\|}{\Delta} + \frac{256n^{-0.5}k\beta^{-0.5}\|E\|^{2}}{\Delta^{2}} \right) \Delta \\ &\leq -2\langle \hat{u}_{1}, \hat{u}_{1,-i}\rangle \left( \hat{u}_{1,-i}^{T}\epsilon_{i} \right) \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*} - \theta_{a}^{*})) - \frac{512\left| \hat{u}_{1,-i}^{T}\epsilon_{i} \right| n^{-1}\|E\|^{2}}{\Delta^{2}} \right\}. \end{split} \right\}. \end{split}$$

From (72) we have

$$\begin{split} \langle \hat{u}_{1,-i}, \hat{u}_1 \rangle &- \frac{512n^{-1} \left\| E \right\|^2}{\Delta^2} \geq 1 - 4 \frac{\left\| E \right\| (n-1)^{-0.5}}{\Delta} - \frac{512n^{-1} \left\| E \right\|^2}{\Delta^2} \\ &\geq 1 - \frac{16n^{-0.5} \left\| E \right\|}{\Delta} \geq \frac{1}{2}, \end{split}$$

assuming  $\frac{\Delta}{n^{-0.5} \|E\|} \ge 64$ . For any  $x, y, z, w \in \mathbb{R}$  such that  $x \ge 0, 1 \ge z \ge 0$ , and  $z |y| > w \ge 0$ , we have  $\mathbb{I}\left\{x \le zy - w\right\} \ge \mathbb{I}\left\{x \le (z - w/|y|) y\right\}$ . We then have,

$$\begin{split} & \mathbb{I}\left\{ \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \hat{\theta}_{a} \right\| \leq \left\| \hat{u}_{1} \hat{u}_{1}^{T} X_{i} - \hat{\theta}_{z_{i}^{*}} \right\| \right\} \\ & \geq \mathbb{I}\left( \left( 1 + \frac{4C_{0}\beta^{-0.5}kn^{-0.5}\|E\|}{\Delta} + \frac{256n^{-0.5}k\beta^{-0.5}\|E\|^{2}}{\Delta^{2}} \right) \Delta \\ & \leq -2\left( 1 - \frac{16n^{-0.5}\|E\|}{\Delta} \right) (\hat{u}_{1,-i}^{T}\epsilon_{i}) \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*} - \theta_{a}^{*})) \right) \\ & \geq \mathbb{I}\left\{ \left( 1 + \frac{C_{1}\beta^{-0.5}n^{-0.5}\|E\|}{\Delta} \right) \Delta \leq -2(\hat{u}_{1,-i}^{T}\epsilon_{i}) \mathrm{sign}(u_{1}^{T}(\theta_{z_{i}^{*}}^{*} - \theta_{a}^{*})) \right\}. \end{split}$$

Since  $\theta_a^* = -\theta_{z_i^*}^*$ , we have  $\operatorname{sign}(u_1^T(\theta_{z_i^*}^* - \theta_a^*)) = \operatorname{sign}(u_1^T\theta_{z_i^*}^*)$ . The proof is complete.  $\Box$ 

PROOF OF THEOREM 3.4. Recall that  $\lambda_1 = \Delta \sqrt{n}/2$ . Same as the proof of Theorem 3.1, we work on the high-probability event (46).

For the upper bound, from Lemma 3.2, there exists some  $\phi \in \Phi$  such that for any  $i \in [n]$ ,  $\mathbb{I}\left\{\hat{z}_i \neq \phi(z_i^*)\right\} \leq \mathbb{I}\left\{\left(1 - C_1\psi_3^{-1}\right)\Delta \leq 2 \left\|\hat{u}_{1,-i}\hat{u}_{-i}^T\epsilon_i\right\|\right\} = \mathbb{I}\left\{\left(1 - C_1\psi_3^{-1}\right)\Delta \leq 2 \left|\hat{u}_{1,-i}^T\epsilon_i\right|\right\}$ , for some  $C_1 > 0$ , where the last inequality is due to that  $\psi_3$  is large. By Davis-Kahan Theorem, we know there exists some  $s_i \in \{-1,1\}$  such that  $\|\hat{u}_{1,-i} - s_iu_1\| \leq 2 \|E\| / (\sqrt{n-1}\Delta) \leq 4\psi_3^{-1}$ . Since  $\langle \hat{u}_{1,-i}, u_1 \rangle \geq 0$  is assumed, we have  $s_i = 1$  for all  $i \in [n]$ . Then

$$\mathbb{I}\left\{\hat{z}_{i} \neq \phi(z_{i}^{*})\right\} \leq \mathbb{I}\left\{\left(1 - C_{1}\psi_{3}^{-1}\right)\Delta \leq 2\left|u_{1}^{T}\epsilon_{i}\right| + 2\left|\left(\hat{u}_{1,-i} - s_{i}u_{1}\right)^{T}\epsilon_{i}\right|\right\} \\ \leq \mathbb{I}\left\{\left(1 - (C_{1} + C_{2})\psi_{3}^{-1}\right)\Delta \leq 2\left|u_{1}^{T}\epsilon_{i}\right|\right\} + \mathbb{I}\left\{C_{2}\psi_{3}^{-1}\Delta \leq 2\left|\left(\hat{u}_{1,-i} - s_{i}u_{1}\right)^{T}\epsilon_{i}\right|\right\},$$

where  $C_2 > 0$  is a constant whose value will be determined later. Due to the independence of  $\hat{u}_{1,-i} - s_i u_1$  and  $\epsilon_i$ , we have  $(\hat{u}_{1,-i} - s_i u_1)^T \epsilon_i \sim \text{SG}(16\psi_3^{-2}\sigma^2)$  and then

$$\mathbb{EI}\left\{C_2\Delta \leq 2\left|\left(\hat{u}_{1,-i} - s_i u_1\right)^T \epsilon_i\right|\right\} \leq 2\exp\left(-\frac{C_2^2 \Delta^2}{128\sigma^2}\right).$$

On the other hand,  $u_1^T \epsilon_i = p^{-\frac{1}{2}} \sum_{j=1}^p \epsilon_{i,j}$  where  $\{\epsilon_{i,j}\}_{j \in [p]}$  are i.i.d. with variance  $\bar{\sigma}^2$ , which can be approximated by a normal distribution. Since the distribution F is sub-Gaussian, its moment generating function exists. Then we can use the following KMT quantile inequality (see Proposition [KMT] of [30]). Let  $Y \stackrel{d}{=} \bar{\sigma}^{-1} p^{-\frac{1}{2}} \sum_{j=1}^p \epsilon_{i,j}$ . There exist some constants  $D, \eta > 0$  and  $Z \sim \mathcal{N}(0, 1)$ , such that whenever  $|Y| \leq \eta \sqrt{p}$ , we have

$$|Y-Z| \leq \frac{DY^2}{\sqrt{p}} + \frac{D}{\sqrt{p}}.$$

Then,

$$\begin{split} &\mathbb{EI}\left\{\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\Delta \le 2\left|u_1^T \epsilon_i\right|\right\} \\ &= \mathbb{EI}\left\{\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \le 2\left|Y\right|\right\} \\ &\le \mathbb{EI}\left\{\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \le 2\left|Z\right| + \frac{2DY^2}{\sqrt{p}} + \frac{2D}{\sqrt{p}}\right\} + \mathbb{EI}\left\{|Y| > \eta\sqrt{p}\right\} \\ &\le \mathbb{EI}\left\{\left(1 - (C_1 + C_2 + C_3 + 2D)\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \le 2\left|Z\right|\right\} + \mathbb{EI}\left\{\frac{2DY^2}{\sqrt{p}} \ge C_3\right\} + \mathbb{EI}\left\{|Y| > \eta\sqrt{p}\right\}, \end{split}$$

where 
$$C_3 > 0$$
 is a constant. Using the fact that  $Y \sim SG(1)$  with zero mean, we have  

$$\mathbb{EI}\left\{\left(1 - (C_1 + C_2)\psi_3^{-1}\right)\Delta \le 2\left|u_1^T\epsilon_i\right|\right\}$$

$$\le 2\exp\left(-\frac{\left(1 - (C_1 + C_2 + C_3 + 2D)\psi_3^{-1}\right)^2\Delta^2}{8\bar{\sigma}^2}\right) + 2\exp\left(-\frac{C_3\sqrt{p}}{4D}\right) + 2\exp\left(-\frac{\eta^2 p}{2}\right)$$

Then we have

$$\begin{split} & \mathbb{E}\ell(\check{z}, z^*) \\ & \leq \frac{1}{n} \sum_{i=1}^n \mathbb{E}\mathbb{I}\left\{ \left(1 - (C_1 + C_2)\psi_3^{-1}\right)\Delta \leq 2\left|u_1^T \epsilon_i\right|\right\} + \frac{1}{n} \sum_{i=1}^n \mathbb{E}\mathbb{I}\left\{C_2\Delta \leq 2\left|(\hat{u}_{1,-i} - s_i u_1)^T \epsilon_i\right|\right\} + e^{-0.5n} \\ & \leq 2\exp\left(-\frac{\left(1 - (C_1 + C_2 + C_3 + 2D)\psi_3^{-1}\right)^2\Delta^2}{8\bar{\sigma}^2}\right) \\ & + 2\exp\left(-\frac{C_2^2\Delta^2}{128\sigma^2}\right) + 2\exp\left(-\frac{C_3\sqrt{p}}{4D}\right) + 2\exp\left(-\frac{\eta^2 p}{2}\right) + e^{-0.5n}, \end{split}$$

where  $e^{-0.5n}$  is the probability that (46) does not hold. Since  $\sigma \leq C\bar{\sigma}$ , when  $C_2$  is chosen to satisfy  $C_2^2/(128C^2) \geq 16$ , we have

$$\mathbb{E}\ell(\check{z}, z^*) \le 2\exp\left(-\frac{\left(1 - C''\psi_3^{-1}\right)^2 \Delta^2}{8\bar{\sigma}^2}\right) + \exp\left(-C''\sqrt{p}\right) + e^{-0.5n},$$

for some constant C'' > 0.

For the lower bound, from (70) we know

$$\mathbb{I}\left\{\check{z}_{i} \neq \phi(z_{i}^{*})\right\} \geq \mathbb{I}\left\{\left(1 + C_{4}\psi_{3}^{-1}\right)\Delta \leq -2(\hat{u}_{1,-i}^{T}\epsilon_{i})\operatorname{sign}(u_{1}^{T}(\theta_{\phi(z_{i}^{*})} - \theta_{3-\phi(z_{i}^{*})}))\right\},\$$

for some constant  $C_4 > 0$  assuming  $\psi_3$  is large. Using the same argument as in the upper bound, we are going to decompose  $\hat{u}_{1,-i}^T \epsilon_i$  into  $u_1^T \epsilon_i$  and  $(\hat{u}_{1,-i} - y_1)^T \epsilon_i$ . Hence,

$$\begin{split} \mathbb{I}\left\{\check{z}_{i} \neq \phi(z_{i}^{*})\right\} &\geq \mathbb{I}\left\{\left(1 + C_{4}\psi_{3}^{-1}\right)\Delta \leq -2(u_{1}^{T}\epsilon_{i})\mathrm{sign}(u_{1}^{T}(\theta_{\phi(z_{i}^{*})} - \theta_{3-\phi(z_{i}^{*})})) - 2\left|(\hat{u}_{1,-i} - s_{i}u_{1})^{T}\epsilon_{i}\right|\right\} \\ &\geq \mathbb{I}\left\{\left(1 + (C_{4} + C_{5})\psi_{3}^{-1}\right)\Delta \leq -2(u_{1}^{T}\epsilon_{i})\mathrm{sign}(u_{1}^{T}(\theta_{\phi(z_{i}^{*})} - \theta_{3-\phi(z_{i}^{*})}))\right\} \\ &- \mathbb{I}\left\{C_{5}\psi_{3}^{-1}\Delta \leq 2\left|(\hat{u}_{1,-i} - s_{i}u_{1})^{T}\epsilon_{i}\right|\right\}, \end{split}$$

for some constant  $C_5 > 0$  whose value to be chosen. Let

$$Y' \stackrel{d}{=} \bar{\sigma}^{-1}(u_1^T \epsilon_i) \operatorname{sign}(u_1^T(\theta_{\phi(z_i^*)} - \theta_{3-\phi(z_i^*)})) = \operatorname{sign}(u_1^T(\theta_{\phi(z_i^*)} - \theta_{3-\phi(z_i^*)})) \bar{\sigma}^{-1} p^{-\frac{1}{2}} \sum_{j=1}^p \epsilon_{i,j}.$$

Then using the same argument above, there exists some  $Z' \sim \mathcal{N}(0,1)$  such that whenever  $Y' \leq \eta' \sqrt{p}$ , we have  $|Y' - Z'| \leq \frac{D'Y'^2}{\sqrt{p}} + \frac{D'}{\sqrt{p}}$  where  $D', \eta' > 0$  are constants. Then

$$\begin{split} &\mathbb{E}\mathbb{I}\left\{\left(1 + (C_4 + C_5)\,\psi_3^{-1}\right)\Delta \leq -2(u_1^T\epsilon_i)\text{sign}(u_1^T(\theta_{\phi(z_i^*)} - \theta_{3-\phi(z_i^*)}))\right\} \\ &= \mathbb{E}\mathbb{I}\left\{\left(1 + (C_4 + C_5)\,\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \leq -2Y'\right\} \\ &\geq \mathbb{E}\mathbb{I}\left\{\left(1 + (C_4 + C_5)\,\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \leq -2Z' - \frac{2DY'^2}{\sqrt{p}} - \frac{2d}{\sqrt{p}}\right\}\mathbb{I}\left\{Y' \leq \eta'\sqrt{p}\right\} \\ &\geq \mathbb{E}\mathbb{I}\left\{\left(1 + (C_4 + C_5 + 2D + C_6)\,\psi_3^{-1}\right)\frac{\Delta}{\bar{\sigma}} \leq -2Z'\right\} - \mathbb{E}\mathbb{I}\left\{\frac{2DY'^2}{\sqrt{p}} \geq C_6\right\} - \mathbb{E}\mathbb{I}\left\{Y' > \eta'\sqrt{p}\right\}. \end{split}$$

$$\mathbb{E}\ell(\check{z}, z^*) \ge 2 \exp\left(-\frac{\left(1 + C'''\psi_3^{-1}\right)^2 \Delta^2}{8\bar{\sigma}^2}\right) - \exp\left(-C'''\sqrt{p}\right) - e^{-0.5n},$$

for some constant C''' > 0.

## D.2. Proofs of Lemma 3.4 and Theorem 3.5.

**PROOF OF LEMMA 3.4.** For the upper bound, we consider the following likelihood ratio test. For any  $x \in \mathbb{R}^p$ , define the two log-likelihood functions as

$$l_1(x) = \sum_{j=1}^p \log f(x_j - \delta)$$
, and  $l_2(x) = \sum_{j=1}^p \log f(x_j + \delta)$ .

Then for each  $i \in [n]$ , define the likelihood ratio test as

$$\hat{z}_i^{\text{LRT}} = \begin{cases} 1, \text{ if } l_1(X_i) \ge l_2(X_i), \\ 2, \text{ otherwise.} \end{cases}$$

Then for any  $i \in [n]$  such that  $z_i^* = 1$ , we have

$$\mathbb{EI}\left\{\hat{z}_{i}^{\mathrm{LRT}}=2\right\} = \mathbb{P}\left(l_{2}(X_{i}) > l_{1}(X_{i})\right) = \mathbb{P}\left(\sum_{j=1}^{p}\log\frac{f(2\delta + \epsilon_{i,j})}{f(\epsilon_{i,j})} > 0\right) = \mathbb{P}\left(\sum_{j=1}^{p}\log\frac{f_{\frac{\Delta}{\sqrt{p}}}(\epsilon_{i,j})}{f_{0}(\epsilon_{i,j})} > 0\right),$$

where we use the fact  $2\delta = \frac{\Delta}{\sqrt{p}}$ . Since  $\Delta$  is a constant, by local asymptotic normality (c.f., Chapter 7, [41]), we have

$$\sum_{j=1}^{p} \log \frac{f_{\frac{\Delta}{\sqrt{p}}}(\epsilon_{i,j})}{f_{0}(\epsilon_{i,j})} \stackrel{d}{\to} \mathcal{N}\left(-\frac{\mathcal{I}\Delta^{2}}{2}, \mathcal{I}\Delta^{2}\right).$$

Then,  $\lim_{p\to\infty} \mathbb{EI}\left\{\hat{z}_i^{\text{LRT}} = 2\right\} \leq C_1 \exp\left(-\mathcal{I}\Delta^2/8\right)$  for some constant  $C_1 > 0$ . We have the same upper bound if  $z_i^* = 2$  instead. Hence,

$$\lim_{p\to\infty} \inf_z \sup_{z^*\in [2]^n} \mathbb{E}\ell(z,z^*) \leq \lim_{p\to\infty} \sup_{z^*\in [2]^n} \mathbb{E}\ell(\hat{z}^{\mathrm{LRT}},z^*) \leq \exp\left(-\frac{\mathcal{I}\Delta^2}{8}\right).$$

For the lower bound, instead of allowing  $z^* \in [2]^n$ , we consider a slightly smaller parameter space. Define  $\mathcal{Z} = \{z \in [2]^n : z_i = 1, \forall 1 \le i \le n/3, z_i = 2, \forall n/3 + 1 \le i \le 2n/3\}.$ Then for any  $z, z' \in \mathcal{Z}$  we have  $\ell(z, z') = n^{-1} \sum_{i=1}^n \mathbb{I}\{z_i \ne z'_i\} \le 1/3$  due to the fact  $n^{-1}\sum_{i=1}^{n} \mathbb{I}\left\{\phi(z_i) \neq z'_i\right\} \ge 1/3$  if  $\phi \neq \text{Id.}$  Hence,

$$\begin{split} \inf_{z} \sup_{z^* \in [2]^n} \mathbb{E}\ell(z, z^*) &\geq \inf_{z} \sup_{z^* \in \mathcal{Z}} \mathbb{E}\ell(z, z^*) \geq n^{-1} \inf_{z} \sup_{z^* \in \mathcal{Z}} \mathbb{E} \sum_{i \in [n]} \mathbb{I} \left\{ z_i \neq z_i^* \right\} \\ &\geq n^{-1} \sum_{i > 2n/3} \inf_{z_i} \sup_{z_i^* \in [2]} \mathbb{E} \mathbb{I} \left\{ z_i \neq z_i^* \right\} = \frac{1}{3} \inf_{z_n} \sup_{z_n^* \in [2]} \mathbb{E} \mathbb{I} \left\{ z_n \neq z_n^* \right\}, \end{split}$$

where it is reduced into a testing problem on whether  $X_n$  has mean  $\theta_1^*$  or  $\theta_2^*$ . According to the Neyman-Pearson Lemma, the optimal procedure is the likelihood ratio test  $\hat{z}_n^{\text{LRT}}$  defined above. By the same argument, we have

$$\liminf_{p \to z} \sup_{z^* \in [2]^n} \mathbb{E}\ell(z, z^*) \ge \frac{1}{3} \liminf_{p \to z_n} \sup_{z_n^* \in [2]} \mathbb{E}\mathbb{I}\left\{z_n \neq z_n^*\right\} \ge C_2 \exp\left(-\frac{\mathcal{I}\Delta^2}{8}\right),$$
  
he constant  $C_2 > 0.$ 

for some constant  $C_2 > 0$ .

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PROOF OF THEOREM 3.5. First, we have the following connection between the Fisher information  $\mathcal{I}$  and the variance  $\bar{\sigma}^2$ :

$$\mathcal{I}\bar{\sigma}^2 = \left(\int \left(\frac{f'}{f}\right)^2 f \mathrm{d}x\right) \left(\int x^2 f \mathrm{d}x\right) \ge \left(\int \frac{f'}{f} x f \mathrm{d}x\right)^2 = \left(\int x f' \mathrm{d}x\right)^2 = 1,$$

where we use Cauchy-Schwarz inequality and the integral by part  $\int xf' dx = \int xf dx - \int f dx = 0 - 1 = -1$ . The equation holds if and only if  $f'/f \propto x$ , which is equivalent to F being normally distributed.

# APPENDIX E: AUXILIARY LEMMAS AND PROPOSITIONS AND THEIR PROOFS

**PROPOSITION E.1.** For Y and  $\hat{Y}$  defined in (1), we have (2) holds assuming  $\sigma_r - \sigma_{r+1} > 2 \| (I - U_r U_r^T) y_n \|$ .

PROOF. Recall the augmented matrix Y' is defined as  $(Y, U_r U_r^T y_n)$ . Note that  $U_r U_r^T Y$  is the best rank-*r* approximation of *Y*. Since

$$\| (I - U_r U_r^T) Y' \|_{\mathrm{F}} = \| ((I - U_r U_r^T) Y, 0) \|_{\mathrm{F}} = \| (I - U_r U_r^T) Y \|_{\mathrm{F}},$$

we have  $U_r U_r^T Y'$  also being the best rank-*r* approximation of Y'. This proves that span $(U_r)$  and  $U_r U_r^T$  are also the leading *r* left singular subspace and projection matrix of Y'. Then  $\hat{U}_r \hat{U}_r^T - U_r U_r^T$  is about the perturbation between  $\hat{Y}$  and Y'.

Let  $\sigma'_r, \sigma'_{r+1}$  be the *r*th and (r+1)th largest singular values of *Y'*, respectively. By Wedin's Thereom (see Section 2.3 of [9]), if  $\sigma'_r - \hat{\sigma}_{r+1} > 0$ , then we have

(73) 
$$\|\sin\Theta(\hat{U}_{r}, U_{r})\|_{\mathrm{F}} \leq \frac{\left\|\hat{Y} - Y'\right\|_{\mathrm{F}}}{\sigma'_{r} - \hat{\sigma}_{r+1}} = \frac{\left\|(I - U_{r}U_{r}^{T})y_{n}\right\|}{\sigma'_{r} - \hat{\sigma}_{r+1}}$$

Regarding the values of  $\sigma'_r$  and  $\sigma'_{r+1}$ , first we have  $\sigma'_r \ge \sigma_r$ . This is because

$$\sigma'_r = \inf_{x \in \operatorname{span}(U_r)} \left\| x^T Y' \right\| = \inf_{x \in \operatorname{span}(U_r)} \left\| \left( x^T Y, x^T y_n \right) \right\| \ge \inf_{x \in \operatorname{span}(U_r)} \left\| x^T Y \right\| \ge \sigma_r.$$

In addition, we have  $\sigma'_{r+1} = \sigma_{r+1}$ , due to the fact that  $(I - U_r U_r^T)Y' = ((I - U_r U_r^T)Y, 0)$ . By Weyl's inequality, we have

$$|\hat{\sigma}_{r+1} - \sigma'_{r+1}| \le ||Y - Y'|| = ||(I - U_r U_r^T)y_n||.$$

Hence, if  $\sigma_r - \sigma_{r+1} > 2 \left\| (I - U_r U_r^T) y_n \right\|$  is further assumed, we have

(74) 
$$\sigma'_{r} - \hat{\sigma}_{r+1} \ge \sigma_{r} - \sigma_{r+1} - \left\| (I - U_{r}U_{r}^{T})y_{n} \right\| \ge \frac{1}{2} \left( \sigma_{r} - \sigma_{r+1} \right).$$

With (73), (74), and the fact  $\|\hat{U}_r\hat{U}_r^T - U_rU_r^T\|_F = \sqrt{2}\|\sin\Theta(\hat{U}_r, U_r)\|_F$  (see Lemma 1 of [9]), the proof is complete.

LEMMA E.1. Let  $E = (\epsilon_1, \ldots, \epsilon_n) \in \mathbb{R}^{p \times n}$  be a random matrix with each column  $\epsilon_i \sim SG_p(\sigma^2), \forall i \in [n]$  independently. Then

$$\mathbb{P}\left(\|E\| \ge 4t\sigma(\sqrt{n} + \sqrt{p})\right) \le \exp\left(-\frac{(t^2 - 3)n}{2}\right).$$

for any  $t \geq 2$ .

PROOF. We follow a standard  $\epsilon$ -net argument. Let  $\mathcal{U}$  and  $\mathcal{V}$  be a 1/4 covering set of the unit sphere in  $\mathbb{R}^p$  and in  $\mathbb{R}^n$ , respectively. That is, for any  $u \in \mathbb{R}^p$  such that ||u|| = 1, there exists a  $u' \in \mathcal{U}$  such that ||u'|| = 1 and  $||u - u'|| \le 1/4$ . Similarly, for any  $v \in \mathbb{R}^n$  such that ||v|| = 1, there exists a  $v' \in \mathcal{V}$  such that ||v'|| = 1 and  $||v - v'|| \le 1/4$ . Then

$$|u^{T}Ev| = |u'^{T}Ev' + u'^{T}E(v - v') + (u - u')^{T}Ev' + (u - u')^{T}E(v - v')|$$
  
$$\leq |u'^{T}Ev'| + |u'^{T}E(v - v')| + |(u - u')^{T}Ev'| + |(u - u')^{T}E(v - v')|$$

Maximizing over u, v on both sides, we have

$$||E|| = \max_{u \in \mathbb{R}^{p}, v \in \mathbb{R}^{n} : ||u|| = ||v|| = 1} |u^{T} E v| \le \max_{u' \in \mathcal{U}, v' \in \mathcal{V}} |u'^{T} E v'| + \frac{1}{4} ||E|| + \frac{1}{4} ||E|| + \frac{1}{16} ||E||.$$

Hence,

$$\|E\| \le 4 \max_{u' \in \mathcal{U}, v' \in \mathcal{V}} \left| u^{'T} E v' \right|.$$

For any  $u' \in \mathcal{U}, v' \in \mathcal{V}$ , we have each  $u'^T \epsilon_i$  being an independent  $SG(\sigma^2)$  and then  $u'^T Ev' \sim SG(\sigma^2)$ . Note  $|U| \leq 9^p \leq e^{3p}$  and similarly  $|V| \leq e^{3n}$ . Then by the tail probability of sub-Gaussian random variable and by the union bound, we have

$$\mathbb{P}\left(\|E\| \le 4t\sigma(\sqrt{n} + \sqrt{p})\right) \le \mathbb{P}\left(\max_{u' \in \mathcal{U}, v' \in \mathcal{V}} \left|u'^T Ev'\right| \le t\sigma(\sqrt{n} + \sqrt{p})\right)$$
$$\le |U| |V| \exp\left(-\frac{t^2\left(\sqrt{n} + \sqrt{p}\right)^2}{2}\right)$$
$$\le \exp\left(-\frac{(t^2 - 3)n}{2}\right),$$

for any  $t \geq 2$ .

LEMMA E.2. Let  $X \sim SG_d(\sigma^2)$ . Consider any  $k \leq d$ . For any matrix  $U = (u_1, \ldots, u_k) \in \mathbb{R}^{d \times k}$  that is independent of X and is with orthogonal columns  $\{u_i\}_{i \in [k]}$ . We have

$$\mathbb{P}\left(\left\|UU^{T}X\right\|^{2} \ge \sigma^{2}(k+2\sqrt{kt}+2t)\right) \le e^{-t}.$$

PROOF. Note that  $tr(UU^T) = tr((UU^T)^2) = k$  and  $||UU^T|| = 1$ . This is a direct consequence of Theorem 1 in [18] for concentration of quadratic forms of sub-Gaussian random vectors.

PROOF OF PROPOSITION 3.1. Define  $\hat{P} = \sum_{i \in [r]} \hat{\lambda}_i \hat{u}_i \hat{v}_i^T$ . Due to the fact that  $\hat{P}$  is the best rank-*r* approximation of *X* in spectral norm and *P* is rank- $\kappa$ , under the assumption that  $\kappa \leq r$ , we have that

$$\left\| \hat{P} - X \right\| \le \|P - X\| = \|E\|.$$

Since  $r \leq k$  is assumed, the rank of  $\hat{P} - P$  his at most 2k, and we have

(75) 
$$\left\| \hat{P} - P \right\|_{\mathrm{F}} \le \sqrt{2k} \left\| \hat{P} - P \right\| \le \sqrt{2k} \left( \left\| \hat{P} - X \right\| + \left\| P - X \right\| \right) \le 2\sqrt{2k} \left\| E \right\|$$

Now, denote  $\hat{\Theta} := (\hat{\theta}_{\hat{z}_1}, \hat{\theta}_{\hat{z}_2}, \dots, \hat{\theta}_{\hat{z}_n})$ . Since  $\hat{\Theta}$  is the solution to the k-means objective (15), we have that

$$\left\| \hat{\boldsymbol{\Theta}} - \hat{\boldsymbol{P}} \right\|_{\mathrm{F}} \leq \left\| \boldsymbol{P} - \hat{\boldsymbol{P}} \right\|_{\mathrm{F}}$$

Hence, by the triangle inequality, we obtain that

$$\left|\hat{\Theta} - P\right\|_{\mathcal{F}} \le 2 \left\|\hat{P} - P\right\|_{\mathcal{F}} \le 4\sqrt{2k} \left\|E\right\|.$$

Now, define the set S as

$$S = \left\{ i \in [n] : \left\| \hat{\theta}_{\hat{z}_i} - \theta^*_{z^*_i} \right\| > \frac{\Delta}{2} \right\}.$$

Since  $\left\{\hat{\theta}_{\hat{z}_i} - \theta_{z_i^*}^*\right\}_{i \in [n]}$  are exactly the columns of  $\hat{\Theta} - P$ , we have that

$$|S| \le \frac{\left\|\hat{\Theta} - P\right\|_{\mathrm{F}}^2}{\left(\Delta/2\right)^2} \le \frac{128k \|E\|^2}{\Delta^2}$$

Under the assumption (16) we have

$$\frac{\beta \Delta^2 n}{k^2 \left\| E \right\|^2} \ge 256$$

which implies

$$|S| \le \frac{\beta n}{2k}$$

We now show that all the data points in  $S^C$  are correctly clustered. We define

$$C_j = \left\{ i \in [n] : z_i^* = j, i \in S^C \right\}, \ j \in [k].$$

The following holds:

- For each j∈ [k], C<sub>j</sub> cannot be empty, as |C<sub>j</sub>| ≥ |{i : z<sub>i</sub>\* = j}| |S| > 0.
  For each pair j, l ∈ [k], j ≠ l, there cannot exist some i ∈ C<sub>j</sub>, i' ∈ C<sub>l</sub> such that ẑ<sub>i</sub> = ẑ<sub>i'</sub>. Otherwise  $\hat{\theta}_{\hat{z}_i} = \hat{\theta}_{\hat{z}_{i'}}$  which would imply

$$\begin{split} \left\| \theta_{j}^{*} - \theta_{l}^{*} \right\| &= \left\| \theta_{z_{i}^{*}}^{*} - \theta_{z_{i'}^{*}}^{*} \right\| \\ &\leq \left\| \theta_{z_{i}^{*}}^{*} - \hat{\theta}_{\hat{z}_{i}} \right\| + \left\| \hat{\theta}_{\hat{z}_{i}} - \hat{\theta}_{\hat{z}_{i'}} \right\| + \left\| \hat{\theta}_{\hat{z}_{i'}} - \theta_{z_{i'}^{*}}^{*} \right\| < \Delta, \end{split}$$

contradicting with the definition of  $\Delta$ .

Since  $\hat{z}_i$  can only take values in [k], we conclude that the sets  $\{\hat{z}_i : i \in C_j\}$  are disjoint for all  $j \in [k]$ . That is, there exists a permutation  $\phi \in \Phi$ , such that

$$\hat{z}_i = \phi(j), \ i \in C_j, \ j \in [k].$$

This implies that  $\sum_{i \in S^C} \mathbb{I}\{\hat{z}_i \neq \phi(z_i^*)\} = 0$ . Hence, we obtain that

$$|\{i \in [n] : \hat{z}_i \neq \phi(z_i^*)\}| \le |S| \le \frac{128k \|E\|^2}{\Delta^2}.$$

Since  $|S| \leq \frac{\beta n}{2k}$  (which means  $\ell(\hat{z}, z^*) \leq \frac{\beta n}{2k}$  from the above display), for any  $\psi \in \Phi$  such that  $\psi \neq \phi$ , we have  $|\{i \in [n] : \hat{z}_i \neq \psi(z_i^*)\}| \geq 2\beta n/k - |S| \geq \beta n/k$ . As a result, we have

$$\ell(\hat{z}, z^*) = \frac{1}{n} |\{i \in [n] : \hat{z}_i \neq \phi(z_i^*)\}| \le \frac{128k \, ||E||^2}{n\Delta^2}.$$

# Moreover, for each $a \in [k]$ , we have

$$\left\|\hat{\theta}_{\phi(a)} - \theta_{a}^{*}\right\|^{2} \leq \frac{\left\|\hat{\Theta} - P\right\|_{\mathrm{F}}^{2}}{\left|\{i \in [n] : \hat{z}_{i} = \phi(a), z_{i}^{*} = a\}\right|} \leq \frac{\left\|\hat{\Theta} - P\right\|_{\mathrm{F}}^{2}}{\frac{\beta n}{k} - |S|} \leq \frac{64k^{2} \left\|E\right\|^{2}}{\beta n}$$